

Foundations for Type-Driven Probabilistic Modelling

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logic-rich & type-rich computation

statistical computation

logic-rich & type-rich computation

- ▶ Expressive type systems: Haskell, OCaml, Rust, Agda, Idris
- ▶ Mechanised mathematics: Agda, Rocq, Isabelle/HOL, Lean
- ▶ Verification: SMT-powered real-world systems

statistical computation

Generative modelling with efficient inference: Monte-Carlo simulation or gradient-based optimisation

This course

Typed interface to probability/statistics

Every concept has:

- ▶ a type
- ▶ associated operations
- ▶ properties in terms of these operations.



course page

Two implementations/models

discrete model

familiar maths
introductory



full model

supports discrete
and
continuous distributions
same language

Motivation: why foundations?

discrete probability

countably supported distributions

good type-structure

(this course)



well-behaved probability

s-finite distributions

over standard Borel spaces



measure theory

standard, established

poor type-structure

continuous probability

Lebesgue measure over \mathbb{R}^n



quasi-Borel spaces

new, experimental

rich type-structure

(this course)

Takeaway

Use types to abstract away from the model

Motivation: why types?

- ▶ **spotlights** meaningful operations

$$\int : (\text{Distribution } X) \times (\text{RandomVariable } X) \rightarrow [0, \infty]$$

- ▶ document **intent**:

probability ($\text{Distribution } X$) vs. density ($X \rightarrow [0, \infty]$) vs. random variable

- ▶ succinctness: omit and elaborate details
- ▶ especially **formal** types, allow using theory correctly without fully understanding it

Part 1: the **discrete** model (now)

- ▶ Motivation
- ▶ Language of probability and distribution
- ▶ Discrete model
- ▶ Simply-typed probability
- ▶ Dependently-typed probability

Part 2: the **full** model

- ▶ Borel sets and measurable spaces
- ▶ Quasi-Borel spaces
- ▶ Type structure & standard Borel spaces
- ▶ Dependently-typed structure
- ▶ Integration



course page



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Language of **probability** & **distribution**

X type (=space) of **values/outcomes**

DX type of **distributions/measures** over X

$PX \subseteq DX$ sub-type of **probability distributions** over X

$\mathcal{B}_X \subseteq \mathcal{P}X$ type of **events**—subsets we wish to measure

\mathbb{W} type of **weights**: values in $[0, \infty]$

\int, \mathbb{E} Lebesgue integration and the expectation operation

Type judgements describe well-formed values/outcomes of a given type, e.g.:

$$\mu : DX, E : \mathcal{B}_X \vdash \text{Ce}_{\mu}[E] : \mathbb{W}$$

(measures weight $\text{Ce}_{\mu}[E]$ of event E according to distribution μ)

Propositions describe properties of well-formed values/outcomes of a given type, e.g.:

$$y_1, y_2 : Y \vdash y_1 \stackrel{Y}{=} y_2 : \text{Prop} \quad \mu : PX, E : \mathcal{B}_X \vdash \text{Pr}_{\mu}[E] = \text{Ce}_{\mu}[E]$$

(probability of event according to probability distribution is its measure)

Empty event

$$\emptyset : \mathcal{B}_X$$

Empty events weight zero

$$\mu : \mathcal{D}X \vdash \text{Ce}_{\mu}[\emptyset] = 0$$

Axioms for events and distributions

Boolean Sub-algebra of Events

$E : \mathcal{B}_X \vdash E^c : \mathcal{B}_X$ $E, F : \mathcal{B}_X \vdash E \cap F : \mathcal{B}_X$ so also: $E, F : \mathcal{B}_X \vdash X, E \cup F : \mathcal{B}_X$

Disjoint additivity

$w, v : \mathbb{W} \vdash w + v : \mathbb{W}$ $E, C : \mathcal{B}_X, \mu : \mathbb{D}X \vdash \text{Ce}_\mu[E] = \text{Ce}_\mu[E \cap C] + \text{Ce}_\mu[E \cap C^c]$

Axioms for events and distributions

Boolean Sub-algebra of Events

$E : \mathcal{B}_X \vdash E^c : \mathcal{B}_X$ $E, F : \mathcal{B}_X \vdash E \cap F : \mathcal{B}_X$ so also: $E, F : \mathcal{B}_X \vdash X, E \cup F : \mathcal{B}_X$

Disjoint additivity

$w, v : \mathbb{W} \vdash w + v : \mathbb{W}$ $E, C : \mathcal{B}_X, \mu : \mathbb{D}X \vdash \text{Ce}_\mu[E] = \text{Ce}_\mu[E \cap C] + \text{Ce}_\mu[E \cap C^c]$

Exercise

Derive 'axiomatically' that:

- ▶ measurement is **monotone**:

$$\mu : \mathbb{D}X, E \subseteq F \vdash \text{Ce}_\mu[E] \leq \text{Ce}_\mu[F]$$

- ▶ the **inclusion-exclusion** principle:

$$\mu : \mathbb{D}X, E, F : \mathcal{B}_X \vdash \text{Ce}_\mu[E \cup F] + \text{Ce}_\mu[E \cap F] = \text{Ce}_\mu[E] + \text{Ce}_\mu[F]$$

Axioms for events and distributions

Consider posets:

$$\omega := (\mathbb{N}, \leq) \quad (\mathcal{B}_X, \subseteq) \quad (\mathbb{W}, \leq)$$

ω -chains in a poset $P = (\underline{P}, \leq)$:

$$P^\omega := \{p_- \in \underline{P}^{\mathbb{N}} \mid p_0 \leq p_1 \leq \dots\}$$

Chain-closure of events and weights

$$E_- : (\mathcal{B}_X, \subseteq)^\omega \vdash \bigcup_n E_n : \mathcal{B}_X \quad w_- : (\mathbb{W}, \leq)^\omega \vdash \sup_n w_n : \mathbb{W}$$

Scott-continuity of measurement

$$E_- : (\mathcal{B}_X, \subseteq)^\omega, \mu : \mathbf{D}X \vdash \mathbf{C}e_\mu [\bigcup_n E_n] = \sup_n \mathbf{C}e_\mu [E_n]$$

Probability distributions have total mass one

$$\mathbf{PX} := \{\mu \in \mathbf{DX} \mid \mathbf{Ce}_\mu[X] = 1\} \quad \mu : \mathbf{PX} \vdash \mathbf{cast} \mu : \mathbf{DX}$$

i.e., if we define:

$$\mathbb{1} := [0,1] \quad \mu : \mathbf{PX}, E : \mathcal{B}_X \vdash \mathbf{Pr}_\mu[E] := \mathbf{Ce}_{\mathbf{cast} \mu}[E] : \mathbb{1}$$

then:

$$\mu : \mathbf{PX} \vdash \mathbf{Pr}_\mu[X] = 1$$

Lebesgue integration w.r.t. a distribution

$$\mu : \mathbb{D}X, f : \mathbb{W}^X \vdash \int \mu(\mathrm{d}x) f(x) : \mathbb{W}$$

(NB: We succinctly write \mathbb{W}^X for the type of functions $X \rightarrow \mathbb{W}$.)

Expectation w.r.t. a probability distribution

$$\mu : \mathbb{P}X, f : \mathbb{W}^X \vdash \mathbb{E}_{x \sim \mu} [f(x)] := \int (\text{cast } \mu)(\mathrm{d}x) f(x) : \mathbb{W}$$

We'll use variations on this notation, e.g.:

$$\int \mathrm{d}\mu f, \int f \mathrm{d}\mu, \int f(x) \mu(\mathrm{d}x), \mathbb{E}_\mu [f]$$

Have: Language and (some) axioms

Want: Model

Today: **discrete** model

Next week: **full** model

Part 1: the **discrete** model (now)

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X : types denote **sets**

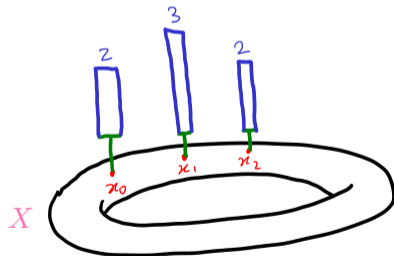
$\mathcal{D}X$: set of **histograms**:

Discrete model

X : types denote **sets**

$\mathcal{D}X$: set of **histograms**:

$$\mathcal{D}X := \{\mu : X \rightarrow \mathbb{W} \mid \mu \text{ is countably supported (next slide)}\}$$



$$\mu x_0 = 2 \quad \mu x_1 = 3 \quad \mu x_2 = 2$$

Support

A subset S **supports** a weight function $\mu : X \rightarrow \mathbb{W}$ when μ is 0 outside S :

$$\mu : \mathbb{W}^X, S : \mathcal{P}X \vdash S \text{ supports } \mu := (\forall x : X. (\mu x > 0) \implies x \in S) : \text{Prop}$$

The subsets supporting a weight function μ are closed under intersections.

\implies There is a smallest supporting subset, called the **support** of μ :

$$\mu : \mathbb{W}^X \vdash \text{supp } \mu := \{x \in X \mid \mu x > 0\}$$

Discrete model

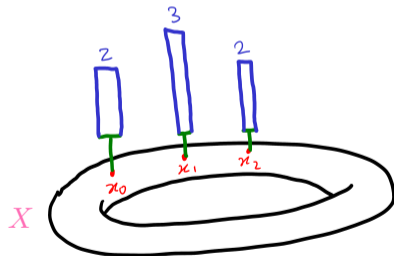
X : types denote **sets**

$\mathcal{D}X$: set of **histograms**:

$$\mathcal{D}X := \{ \mu : X \rightarrow \mathbb{W} \mid \mu \text{ is countably supported} \}$$

$$:= \{ \mu : X \rightarrow \mathbb{W} \mid \exists S \in \mathcal{P}X. S \text{ is countable} \}$$

$$:= \{ \mu : X \rightarrow \mathbb{W} \mid \text{supp } \mu \text{ is countable} \}$$



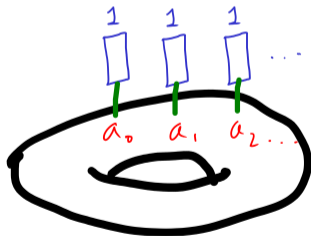
$$\mu x_0 = 2 \quad \mu x_1 = 3 \quad \mu x_2 = 2$$

Example distributions

Counting distribution

Counts the outcomes in a countable subset:

$$S : \mathcal{P}_{\text{ctbl}} X \vdash \#_S := \left(\lambda x. \begin{cases} x \in S : 1 \\ x \notin S : 0 \end{cases} \right) : \mathbb{D}X$$

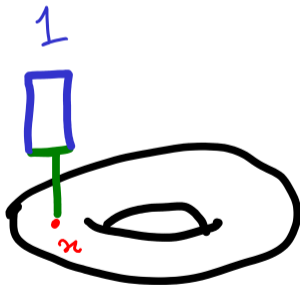


Example distributions

Dirac

A point mass:

$$x : X \vdash \delta_x := \left(\lambda x'. \begin{cases} x' = x : 1 \\ x' \neq x : 0 \end{cases} \right) : DX$$



(NB: $x : X \vdash \delta_x = \#_{\{x\}}$.)

Zero

No mass anywhere:

$$\vdash \mathbf{0} := \underline{0} := (\lambda x.0) : DX$$

(NB: $\vdash \mathbf{0} = \#\emptyset$.)

Discrete model

X : types denote **sets**

\mathbf{DX} : set of **histograms**:

$$\mathbf{DX} := \{ \mu : X \rightarrow \mathbb{W} \mid \mu \text{ is countably supported} \}$$

\mathcal{B}_X : **every subset** can be measured:

$$\mathcal{B}_X := \mathcal{P}X$$

Measurement: weighted sum of all (supported) outcomes:

$$\begin{aligned} \mu : \mathbf{DX}, E : \mathcal{B}_X \vdash \mathbf{Ce}_\mu[E] &:= \sum_{x \in E} \mu x \\ &:= \sum_{x \in E \cap \text{supp } \mu} \mu x \end{aligned}$$

$$\text{NB: } \mu : \mathbf{DX}, E : \mathcal{B}_X, S : \mathcal{P}_{\text{ctbl}}X, S \text{ supports } \mu \vdash \mathbf{Ce}_\mu[E] = \sum_{x \in E \cap S} \mu x.$$

Example measurements

(NB: $\mu : \mathcal{D}X$, $E : \mathcal{B}_X$, $S : \mathcal{P}_{\text{ctbl}}X$, S supports $\mu \vdash \text{Ce}_\mu[E] = \sum_{x \in E \cap S} \mu x$.)

Counting distribution

counts supported outcomes

$$S : \mathcal{P}_{\text{ctbl}}X, E : \mathcal{B}_X \vdash \underset{\#s}{\text{Ce}}[E] = |E \cap S| := \begin{cases} E \cap S \text{ has } n \in \mathbb{N} \text{ elements:} & n \\ E \cap S \text{ is infinite:} & \infty \end{cases}$$

Example measurements

(NB: $\mu : \mathcal{D}X$, $E : \mathcal{B}_X$, $S : \mathcal{P}_{\text{ctbl}}X$, S supports $\mu \vdash \text{Ce}_\mu[E] = \sum_{x \in E \cap S} \mu x$.)

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Dirac

detects given outcome:

$$x : X, E : \mathcal{B}_X \vdash \text{Ce}_{\delta_x}[E] = \begin{cases} x \in E : & 1 \\ x \notin E : & 0 \end{cases}$$

Example measurements

(NB: $\mu : \mathcal{D}X$, $E : \mathcal{B}_X$, $S : \mathcal{P}_{\text{ctbl}}X$, S supports $\mu \vdash \text{Ce}_\mu[E] = \sum_{x \in E \cap S} \mu x$.)

Counting distribution

counts supported outcomes

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Dirac

detects given outcome:

$$x : X, E : \mathcal{B}_X \vdash \text{Ce}_{\delta_x}[E] = \begin{cases} x \in E : & 1 \\ x \notin E : & 0 \end{cases}$$

Zero

measures every event as zero:

$$E : \mathcal{B}_X \vdash \text{Ce}_0[E] = 0$$

Exercise

$$\mu : \mathbf{D} \quad \vdash \text{Ce}_{\mu}[\emptyset] = 0$$

$$E, C : \mathcal{B}_X, \mu : \mathbf{D} \quad \vdash \text{Ce}_{\mu}[E] = \text{Ce}_{\mu}[E \cap C] + \text{Ce}_{\mu}[E \cap C^c]$$

$$E_{-} : (\mathcal{B}_X, \subseteq)^{\omega}, \mu : \mathbf{D} \quad \vdash \text{Ce}_{\mu}\left[\bigcup_n E_n\right] = \sup_n \text{Ce}_{\mu}[E_n]$$

Kernel

$k : X \rightsquigarrow Y$ from X to Y : function $k : X \rightarrow \mathbf{D}Y$.

Kernels are open/parameterised distributions.

Examples

Dirac and the counting distribution form kernels:

$$\delta_- : X \rightsquigarrow \mathbf{D}X \quad \#_- : \mathcal{P}_{\text{ctbl}} X \rightsquigarrow \mathbf{D}X$$

NB: This definition is **internal**: when we consider the full model, we will define kernels as those functions internal to the model rather than the set-theoretic functions.

Kock integral

$$\mu : \mathbf{D}X, k : (\mathbf{D}Y)^X \vdash \oint d\mu k : \mathbf{D}Y$$

This **distribution-valued** integral is implicit in many probability texts. It corresponds to integrating against an arbitrary weight function or random variable.

Discrete model interpretation

$$\begin{aligned} \oint d\mu k &:= \lambda y. \sum_{x \in X} \mu x \cdot k(x; y) \\ &:= \lambda y. \sum_{x \in \text{supp } \mu} \mu x \cdot k(x; y) \end{aligned}$$

NB1: we write $k(x; y) := k(x)(y)$ for the uncurried function.

NB2: $\mu : \mathbf{D}X, k : (\mathbf{D}Y)^X, S : \mathcal{P}_{\text{ctbl}} X, S \text{ supports } \mu \vdash \oint d\mu k = \lambda y. \sum_{x \in S} \mu x \cdot k(x; y)$

Weak Disintegration Problem (non-standard terminology)

Input: distributions $\mu : \mathbf{D}\Theta$, $\nu : \mathbf{D}X$

Output: kernel $k : \Theta \rightsquigarrow X$ such that: $\nu = \int d\mu k$.

Such a **weak disintegration** of ν w.r.t. μ provides an ‘explanation’ of an observed distribution $\nu \in \mathbf{D}X$ in terms of a given distribution on parameters $\mu \in \mathbf{D}\Theta$. I use the term ‘explanation’ because it explains how the parameters transform into observations.

Example

Weak Disintegration Problem (non-standard terminology)

Input: distributions $\mu : \mathbf{D}\Theta$, $\nu : \mathbf{D}X$

Output: kernel $k : \Theta \rightsquigarrow X$ such that: $\nu = \oint d\mu k$.

Example disintegration

For $n \in \mathbb{N}$, write $\mathbf{Fin} n := \{0, \dots, n-1\}$. For countable X , write $\# := \#_X : \mathbf{D}X$.

Here is a disintegration of $\# \in \mathbf{D}((\mathbf{Fin} 2)^{\mathbf{Fin}(n+1)})$ w.r.t. $\# \in \mathbf{D}(\mathbf{Fin} 2)$:

$$k(x; f) := \begin{cases} fn = x : & 1 \\ \text{otherwise:} & 0 \end{cases} \quad \text{Indeed: } \left(\oint d\# k \right) f = \sum_{b \in \mathbf{Fin} 2} \overbrace{\# b}^1 \cdot k(b; f) = k(0; f) + k(1; f)$$

$f : \mathbf{Fin}(n+1) \rightarrow \mathbf{Fin} 2$ function
so can take only one value: 0 or 1

$$\downarrow \\ = 1 = \# f$$

Sub-type of probability distributions

Sub-types

Given type X and $x : X \vdash \varphi : \mathbf{Prop}$, take the **sub-type** and the **coercion** as follows:

$$\{x : X \mid \varphi\} \subseteq X \quad y : \{x : X \mid \varphi\} \vdash \mathbf{cast} \, y := y : X$$

we **lift** values in X that satisfy φ to the sub-type:

$$\frac{\Gamma \vdash M : X \quad \Gamma \vdash \varphi [x \mapsto M]}{\Gamma \vdash \mathbf{lift} M : \{x : X \mid \varphi\}} \quad \frac{\Gamma \vdash M : X \quad \Gamma \vdash \{\varphi\} x \mapsto M}{\Gamma \vdash \mathbf{cast}(\mathbf{lift} M) = M}$$

The axiom implies that $\mathbf{lift} M$ lifts M along \mathbf{cast} . Moreover:

$$y : \{x \in X \mid \varphi\} \vdash \mathbf{lift}(\mathbf{cast} \, y) = y \quad y : \{x \in X \mid \varphi\} \vdash \varphi [x \mapsto \mathbf{cast} \, y]$$

i.e., the lifting is unique and elements in the sub-type satisfy φ .

Sub-type of probability distributions

Magnitude and probability distributions

$$\mu : DX \vdash \|\mu\| := \underset{\mu}{\text{Ce}} [X] : W \quad PX := \{\mu \in DX \mid \|\mu\| = 1\} \quad \mathbb{I} := [0,1] := \{w \in W \mid w \leq 1\}$$

Event probability

$$\mu : PX, E : \mathcal{B}_X \vdash \underset{\mu}{\text{Pr}} [E] := \text{lift} \left(\underset{\text{cast } \mu}{\text{Ce}} [E] \right) : \mathbb{I}$$

Stochastic kernel

$k : X \rightsquigarrow Y$ from X to Y : function $X \rightarrow PY$.

NB: in the **discrete model** these distinctions and rules amount to pure pedantry. This pedantry will pay off in the **full model**.

Lifting Dirac and Kock

Lemma

Dirac kernels $\delta_- : X \rightarrow DX$ lift along cast :

$$x : X \vdash \|\delta_x\| = \underset{\delta_x}{\text{Ce}}[X] = 1 \quad \text{so we can overload:}$$
$$\begin{array}{ccc} & \delta_- & PX \\ X & \begin{array}{c} \dashrightarrow \\ \rightarrow \end{array} & \\ & \delta_- & DX \\ & & \downarrow \text{cast} \\ & & \end{array} \quad =$$

Kock integrals of stochastic kernels by probability distributions lift along cast :

$$\mu : PX, k : (PY)^X \vdash \text{Ce}_{\oint(\text{cast } \mu)(dx) \text{ cast}(kx)}[Y] = 1$$

so we can overload:

$$\begin{array}{ccc} (PX) \times (PY)^X & \begin{array}{c} \dashrightarrow \\ \rightarrow \end{array} & PY \\ \text{cast} \times (\text{cast} \circ) \downarrow & = & \downarrow \text{cast} \\ (DX) \times (DY)^X & \xrightarrow{\oint} & DY \end{array}$$

Proposition

The triple $(\mathbf{D}, \delta_-, \oint)$ forms a monad over **Set**:

$$\begin{array}{l}
 x : X, k : (\mathbf{D}Y)^X \quad \vdash \oint d\delta_x k = k x \\
 \mu : \mathbf{D}X \quad \vdash \oint \mu(dx)\delta_x = \mu \\
 \mu : \mathbf{D}X, k : (\mathbf{D}Y)^X, \ell : (\mathbf{D}Z)^Y \quad \vdash \oint (\oint \mu(dx)k x) (dy)\ell y = \oint \mu(dx) \oint k(x; dy)\ell y
 \end{array}$$

Corollary

The triple $(\mathbf{P}, \delta_-, \oint)$ forms a monad over **Set**.

Weighted average

Lebesgue integral

Integration is the raison d'être for distributions:

$$\mu : \mathbf{D}X, f : \mathbb{W}^X \vdash \int d\mu f : \mathbb{W}$$

In the **discrete model**:

$$\int d\mu f := \sum_{x \in X} (\mu x) \cdot (f x) := \sum_{x \in \text{supp } \mu} (\mu x) \cdot (f x)$$

As usual, replace $\text{supp } \mu$ by any countable supporting set:

$$\mu : \mathbf{D}X, f : \mathbb{W}^X, S : \mathcal{P}X, S \text{ supports } \mu \vdash \int d\mu f = \sum_{x \in S} (\mu x) \cdot (f x)$$

Expectation

To emphasise that some μ is a probability distribution, we will use the notation:

$$\mu : \mathbf{P}X, f : \mathbb{W}^X \vdash \mathbb{E}_\mu[f] := \int d(\text{cast } \mu) f : \mathbb{W}$$

When calculating, however, we will usually use \int and implicitly `cast` any probability distribution to its corresponding distribution.

Boolean type

The simplest kind of distinguishing outcomes:

$$\mathbb{B} := \{\mathbf{True}, \mathbf{False}\} \quad \frac{\Gamma \vdash M : \mathbb{B} \quad \Gamma \vdash N_1 : X \quad \Gamma \vdash N_2 : X}{\Gamma \vdash \text{if } M \text{ then } N_1 \text{ else } N_2 : X}$$

Iverson bracket

Lets us replace Boolean propositions with arithmetic expressions:

$$b : \mathbb{B} \vdash [b] := (\text{if } b \text{ then } 1 \text{ else } 0) : \mathbb{W}$$

For example:

$$b : \mathbb{B}, w, v : \mathbb{W} \vdash \text{if } b \text{ then } w \text{ else } v = [b] \cdot w + (1 - [b]) \cdot w$$

Bernoulli kernel

Single trial succeeding with the given probability:

$$\mathbf{B} : \mathbb{I} \rightsquigarrow \mathbb{B} \quad \mathbf{B}p := \lambda b. \begin{cases} b = \mathbf{True} : & p \\ b = \mathbf{False} : & 1 - p \end{cases}$$

For example, for a payoff of 10 units if the trial succeeds then the expected payoff is:

$$\mathbb{E}_{b \sim \mathbf{B} \frac{1}{4}} [[b] \cdot 10] = \frac{1}{4} \cdot 10 + (1 - \frac{1}{4}) \cdot 0 = \frac{10}{4} + 0 = \frac{5}{2}$$

Proposition

Membership testing induces an isomorphism between events and Boolean propositions:

$$(\in) : \mathcal{B}_X \xrightarrow{\cong} \mathbb{B}^X$$

Its inverse sends each Boolean property to the set of outcomes satisfying it:

$$\frac{x : X \vdash M : \mathbb{B}}{\{x \in X \mid M\} : \mathcal{B}_X} \quad \{x \in X \mid \varphi x\} := \{x \in X \mid \varphi x = \mathbf{True}\}$$

Characteristic function

represents an event as weight functions: $E : \mathcal{B}_X \vdash [- \in E] : \mathbb{W}^X$

By the above proposition, every (internal) $\{0, 1\}$ -valued weight function is the characteristic function of some event, namely, the inverse image of 1 .

Measurement through integration

Lemma

We can replace event measurement by integration of characteristic functions:

$$\mu : \mathbf{DX}, E : \mathcal{B}_X \vdash \mathbf{Ce}_{\mu}[E] = \int \mu(dx) [x \in E]$$

We can deduce properties for $\mathbf{Ce}[-]$ and $\mathbf{Pr}[-]$ from those of the Lebesgue integral.

Notation:

$$\frac{\Gamma \vdash \mu : \mathbf{DX} \quad \Gamma, x : X \vdash M : \mathbb{B}}{\Gamma \vdash \mathbf{Ce}_{x \sim \mu}[M] := \mathbf{Ce}_{\mu}[\{x \in X | M\}] : \mathbb{W}}$$

and similarly for $\mathbf{Pr}_{x \sim \mu}[M]$.

Language of **probability** & **distribution** (recap)

X type of **values/outcomes**

DX type of **distributions/measures** over X

$PX \subseteq DX$ sub-type of **probability distributions** over X

$\mathcal{B}_X \subseteq \mathcal{P}X$ type of **events**—subsets we wish to measure

\mathbb{W} type of **weights**: values in $[0, \infty]$

\int, \mathbb{E} Lebesgue integration and the expectation operation

Type judgements describe well-formed values/outcomes of a given type, e.g.:

$$\mu : DX, E : \mathcal{B}_X \vdash \text{Ce}_{\mu}[E] : \mathbb{W}$$

(measures weight $\text{Ce}_{\mu}[E]$ of event E according to distribution μ)

Propositions describe properties of well-formed values/outcomes of a given type, e.g.:

$$y_1, y_2 : Y \vdash y_1 = y_2 : \text{Prop} \quad \mu : PX, E : \mathcal{B}_X \vdash \text{cast}_{\mu} \text{Pr}[E] = \text{Ce}_{\mu}[E]$$

(probability of event according to probability distribution is its measure)

Part 1: the **discrete** model (now)

- ▶ Motivation
- ▶ Language of probability and distribution
- ▶ Discrete model
- ▶ **Simply-typed probability**
- ▶ Dependently-typed probability

Part 2: the **full** model

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Compositional building blocks for modelling

- ▶ Affine combinations of distributions
- ▶ Product measures (\otimes) : $\mathbf{D}X \times \mathbf{D}Y \rightarrow \mathbf{D}(X \times Y)$
- ▶ Random elements and their laws (push-forward measure):
 $(\lambda(\mu, \alpha) \cdot \mu_\alpha) : \mathbf{D}\Omega \times X^\Omega \rightarrow \mathbf{D}X$

NB:

Standard vocabulary

- ▶ Joint and marginal distributions
- ▶ Independence
- ▶ Distribution/probability preservation and invariance
- ▶ Density and absolute continuity
- ▶ Almost certain/sure properties

- ▶ Dirac kernel $\delta_- : X \rightarrow \mathbf{D}X$
- ▶ Kock integration
 $\oint : \mathbf{D}X \times (\mathbf{D}Y)^{\mathbf{D}X} \rightarrow \mathbf{D}Y$

Compositional building blocks for modelling

- ▶ Affine combinations of distributions
- ▶ Product measures $(\otimes) : \mathbf{D}X \times \mathbf{D}Y \rightarrow \mathbf{D}(X \times Y)$
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Affine combinations of distributions: scaling

Scaling distributions

$$w : \mathbb{W}, \mu : \mathbf{DX} \vdash w \cdot \mu : \mathbf{DX}$$

In the discrete model:

$$w \cdot \mu := \lambda x. w \cdot \mu x \quad \text{supp}(w \cdot \mu) \subseteq \text{supp } \mu$$

The function $(\cdot) : \mathbb{W} \times \mathbf{DX} \rightarrow \mathbf{DX}$ is a **monoid action** for the monoid $(\mathbb{W}, (\cdot), \mathbf{1})$:

$$\mu : \mathbf{DX} \vdash \mathbf{1} \cdot \mu = \mu \quad w, v : \mathbb{W}, \mu : \mathbf{DX} \vdash w \cdot (v \cdot \mu) = (w \cdot v) \cdot \mu$$

Integration and measurement are homogeneous w.r.t. scaling:

$$w : \mathbb{W}, \mu : \mathbf{DX}, k : (\mathbf{DY})^X \vdash \oint d(w \cdot \mu)k = w \cdot \oint d\mu k$$

$$w : \mathbb{W}, \mu : \mathbf{DX}, f : \mathbb{W}^X \vdash \int d(w \cdot \mu)f = w \cdot \int d\mu f$$

$$w : \mathbb{W}, \mu : \mathbf{DX}, E : \mathcal{B}_X \vdash \text{Ce}_{w \cdot \mu}[f] = w \cdot \text{Ce}_{\mu}[f]$$

Affine combinations of distributions: scaling

Normalisation

$$\mu : \mathbf{D}X, \|\mu\| \neq 0, \infty \vdash \frac{\mu}{\|\mu\|} := \text{lift} \left(\frac{1}{\|\mu\|} \cdot \mu \right) : \mathbf{P}X$$

measurement is homogeneous

$$\text{Indeed: } \left\| \frac{\mu}{\|\mu\|} \right\| = \left\| \frac{1}{\|\mu\|} \cdot \mu \right\| \stackrel{\downarrow}{=} \frac{1}{\|\mu\|} \cdot \|\mu\| = 1$$

Discrete uniform / categorical distribution

Random unbiased choice between finitely many options/categories:

$$S : \mathcal{P}_{\text{fin}}(X), S \neq \emptyset \vdash \mathbf{U}_S := \frac{\text{lift}\#_S}{\|\text{lift}\#_S\|} : \mathbf{P}X$$

In the discrete model:

$$\mathbf{U}_S = \lambda x. \begin{cases} x \in S : \frac{1}{|S|} \\ x \notin S : 0 \end{cases}$$

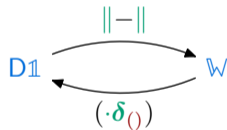
so: $x : X \vdash \mathbf{U}_{\{x\}} = \delta_x$.

Unit type

$$\mathbb{1} := \{()\}$$

Proposition

The following two functions are mutually inverse:



Proof

Calculate: $\mu : \mathbb{D}\mathbb{1} \vdash \mu \mapsto \mu () \mapsto \lambda().\mu () = \mu$ and $w : \mathbb{W} \vdash w \mapsto \lambda().w \mapsto w$. ■

Proposition

We can recover Lebesgue integration from Kock integration:

$$\begin{array}{ccc} DX \times W^X & \xrightarrow{\text{id} \times (\cong \circ)} & DX \times (D1)^X \\ \int \downarrow & = & \downarrow \oint \\ W & \xleftarrow{\cong} & D1 \end{array}$$

Since measurement also reduced to Lebesgue integration, it usually suffices to prove properties of Kock integration and derive them for Lebesgue integration and for measurement.

Affine combinations of distributions: addition

Summation

$$\mu_- : (\mathbf{DX})^I, I \text{ countable} \vdash \sum_{i \in I} \mu_i : \mathbf{DX}$$

In the discrete model:

$$\sum_{i \in I} \mu_i := \lambda x. \sum_{i \in I} \mu_i x \quad \text{supp } \sum_{i \in I} \mu_i = \bigcup_{i \in I} \text{supp } \mu_i$$

Affine and convex combinations

An **affine** combination is a countable sequence of weights $w_- : \mathbb{W}^I$.

It is **convex** when $\sum_{i \in I} w_i = 1$.

Bernoulli revisited

We can express the Bernoulli distribution as follows:

$$p : \mathbb{1} \vdash \mathbf{B}p = \text{lift } (p \cdot \delta_{\mathbf{True}} + (1 - p) \cdot \delta_{\mathbf{False}}) : \mathbf{PB}$$

Affinity of integration and convexity of expectation

Theorem (Multi-linearity)

The Kock and Lebesgue integrals and measurement are affine in each argument:

$$\mu_- : (\mathbf{D}X)^I, w_- : \mathbb{W}^I, k : X \rightsquigarrow Y \vdash \oint d \left(\sum_{i \in I} w_i \cdot \mu_i \right) k = \sum_{i \in I} w_i \cdot \oint d \mu_i k$$

$$\mu : \mathbf{D}X, w_- : \mathbb{W}^I, k_- : (X \rightsquigarrow B)^I \vdash \oint d \mu \left(\sum_{i \in I} w_i \cdot k_i \right) = \sum_{i \in I} w_i \cdot \oint d \mu k_i$$

$$\mu_- : (\mathbf{D}X)^I, w_- : \mathbb{W}^I, \varphi : \mathbb{W}^X \vdash \int d \left(\sum_{i \in I} w_i \cdot \mu_i \right) \varphi = \sum_{i \in I} w_i \cdot \int d \mu_i \varphi$$

$$\mu : \mathbf{D}X, w_- : \mathbb{W}^I, \varphi_- : (\mathbb{W}^X)^I \vdash \int d \mu \left(\sum_{i \in I} w_i \cdot \varphi_i \right) = \sum_{i \in I} w_i \cdot \int d \mu \varphi_i$$

$$\mu_- : (\mathbf{D}X)^I, w_- : \mathbb{W}^I, E : \mathcal{B}_X \vdash \sum_{i \in I} \text{Ce}_{w_i \cdot \mu_i} [E] = \sum_{i \in I} w_i \cdot \text{Ce}_{\mu_i} [E]$$

This theorem, a working horse in probability, has several important consequences:

Proposition

The isomorphism $D\mathbb{1} \cong \mathbb{W}$ is a σ -semiring isomorphism:

$$\left(D\mathbb{1}, \sum, (\cdot) \right) \cong \left(\mathbb{W}, \sum, (\cdot) \right)$$

and $(\cdot) : \mathbb{W} \times DX \rightarrow DX$ makes each DX into a \mathbb{W} -module:

$$\left(\sum_{i \in I} w_i \right) \cdot \mu = \sum_{i \in I} (w_i \cdot \mu) \quad w \cdot \sum_{i \in I} \mu_i = \sum_{i \in I} w \cdot \mu_i$$

Convex combinations of probability distributions

Lemma

Convex combination lifts to probability distributions:

$$w_- : \mathbb{W}^I, \mu_- : (\mathbf{P}X)^I, I \text{ countable}, \sum_{i \in I} w_i = 1 \vdash$$

$$\sum_{i \in I} w_i \cdot \mu_i := \text{lift} \sum_{i \in I} w_i \cdot (\text{cast } \mu_i) : \mathbf{P}X$$

Proof

Calculate: $\left\| \sum_{i \in I} w_i \cdot (\text{cast } \mu_i) \right\| = \sum_{i \in I} w_i \cdot \|\text{cast } \mu_i\| = \sum_{i \in I} w_i \cdot 1 = 1$ ■

Convex combinations of probability distributions

Corollary (Multi-convexity)

Stochastic Kock integration, expectation and measurement are convex:

$$\mu_- : (\mathbf{D}X)^I, w_- : \mathbb{W}^I, k : X \rightsquigarrow Y, \sum_{i \in I} w_i = 1 \vdash \oint d \left(\sum_{i \in I} w_i \cdot \mu_i \right) k = \sum_{i \in I} w_i \cdot \oint d\mu_i k$$

$$\mu : \mathbf{D}X, w_- : \mathbb{W}^I, k_- : (X \rightsquigarrow B)^I, \sum_{i \in I} w_i = 1 \vdash \oint d\mu \left(\sum_{i \in I} w_i \cdot k_i \right) = \sum_{i \in I} w_i \cdot \oint d\mu k_i$$

$$\mu_- : (\mathbf{D}X)^I, w_- : \mathbb{W}^I, \varphi : \mathbb{W}^X, \sum_{i \in I} w_i = 1 \vdash \mathbb{E}_{\sum_{i \in I} w_i \cdot \mu_i} [\varphi] = \sum_{i \in I} w_i \cdot \mathbb{E}_{\mu_i} [\varphi]$$

$$\mu : \mathbf{D}X, w_- : \mathbb{W}^I, \varphi_- : (\mathbb{W}^X)^I, \sum_{i \in I} w_i = 1 \vdash \mathbb{E}_\mu \left[\sum_{i \in I} w_i \cdot \varphi_i \right] = \sum_{i \in I} w_i \cdot \mathbb{E}_\mu [\varphi_i]$$

$$\mu_- : (\mathbf{D}X)^I, w_- : \mathbb{W}^I, E : \mathcal{B}_X, \sum_{i \in I} w_i = 1 \vdash \sum_{i \in I} \Pr_{w_i \cdot \mu_i} [E] = \sum_{i \in I} w_i \cdot \Pr_{\mu_i} [E]$$

Product distribution

$$\mu : \mathbf{D}X, \nu : \mathbf{D}Y \vdash \mu \otimes \nu := \int \mu(\mathrm{d}x) \int \nu(\mathrm{d}y) \delta_{(x,y)} : \mathbf{D}(X \times Y)$$

In the discrete model:

$$\mu \otimes \nu = \lambda(x, y) \cdot (\mu x) \cdot (\nu y) \quad \text{supp}(\mu \otimes \nu) = (\text{supp } \mu) \times (\text{supp } \nu)$$

Example: counting distribution on product space

$$S : \mathcal{P}_{\text{fin}}(X), T : \mathcal{P}_{\text{fin}}(Y) \vdash \#_{S \times T} \stackrel{\mathbf{D}(X \times Y)}{=} \#_S \otimes \#_T$$

Indeed: $\text{supp}(\#_S \otimes \#_T) = S \times T = \text{supp } \#_{S \times T}$ and for $(x, y) \in S \times T$:

$$(\#_S \otimes \#_T)(x, y) = 1 \cdot 1 = 1 = \#_{S \times T}(x, y)$$

Notation:
$$\frac{\Gamma \vdash M : \mathbf{D}(X \times Y) \quad \Gamma, x : X, y : Y \vdash K : \mathbf{D}Z}{\Gamma \vdash \oint M(\mathrm{d}x, \mathrm{d}y)K := \oint \mathrm{d}M(\lambda(x, y).K) : \mathbf{D}Z}$$

Theorem (Fubini-Tonelli)

We can integrate products in any order:

$$\mu : \mathbf{D}X, \nu : \mathbf{D}Y, k : (\mathbf{D}Z)^{X \times Y} \vdash$$

$$\oint \mu(\mathrm{d}x) \oint \nu(\mathrm{d}y) k(x, y) = \oint (\mu \otimes \nu)(\mathrm{d}x, \mathrm{d}y) k(x, y) = \oint \nu(\mathrm{d}y) \oint \mu(\mathrm{d}x) k(x, y)$$

$$\mu : \mathbf{D}X, \nu : \mathbf{D}Y, \varphi : \mathbf{W}^{X \times Y} \vdash$$

$$\int \mu(\mathrm{d}x) \int \nu(\mathrm{d}y) \varphi(x, y) = \iint (\mu \otimes \nu)(\mathrm{d}x, \mathrm{d}y) \varphi(x, y) = \int \nu(\mathrm{d}y) \int \mu(\mathrm{d}x) \varphi(x, y)$$

Applying Fubini-Tonelli

Theorem (Rule of Product)

We can factor out products:

$$\begin{aligned} \mu : \mathbf{D}X, f : \mathbb{W}^X, \nu : \mathbf{D}Y, g : \mathbb{W}^Y \vdash & \iint (\mu \otimes \nu)(dx, dy) f x \cdot g y = \left(\int d\mu f \right) \cdot \left(\int d\nu g \right) \\ \mu : \mathbf{D}X, E : \mathcal{B}_X, \nu : \mathbf{D}Y, F : \mathcal{B}_Y \vdash & \text{Ce}_{\mu \otimes \nu} [E \times F] = \text{Ce}_{\mu} [E] \cdot \text{Ce}_{\nu} [F] \end{aligned}$$

Theorem

The product lifts to probability distributions:

$$\mu : \mathbf{P}X, \nu : \mathbf{P}Y \vdash (\mu \otimes \nu) := \text{lift}(\text{cast } \mu \otimes \text{cast } \nu) : \mathbf{P}(X \times Y)$$

Binomial distribution

the number of successful outcomes of n independent Bernoulli trials:

$$\mathbf{B}_n : \mathbb{1} \rightsquigarrow \mathbf{P}(\mathbf{Fin}(1+n)) \quad \mathbf{B}_0 p := \delta_0 : \mathbf{P}(\mathbf{Fin} 1)$$
$$\mathbf{B}_{1+n} p := \iint (\mathbf{B}_n p \otimes \mathbf{B} p)(dc, db) (\text{if } b \text{ then } \delta_{1+c} \text{ else } \delta_c) : \mathbf{P}(\mathbf{Fin}(2+n))$$

We can prove by induction on n , using Fubini-Tonelli and the Iverson bracket that:

$$p : \mathbb{1}, k : \mathbf{Fin}(1+n) \vdash \Pr_{c \sim \mathbf{B}_n p} [c = k] = \binom{n}{k}$$

Push-forward distributions

Random element

in X any (internal) function:

$$\mu : \mathbf{D}\Omega \vdash \alpha : \Omega \rightarrow X$$

Law

of a random element is the distribution:

$$\mu : \mathbf{D}\Omega, \alpha : X^\Omega \vdash \mu_\alpha := \int \mu(d\omega) \delta_{\alpha\omega} : \mathbf{D}X$$

Example

Represent outcomes of die roll by $\mathbf{D}6 := \{1, 2, \dots, 6\}$, and two rolls by $\mathbf{D}6 \times \mathbf{D}6$.

The sum of the rolls is a random element:

$$(+): \mathbf{D}6 \times \mathbf{D}6 \rightarrow \mathbb{N}$$

The law of the distribution $\# \otimes \#$ counts the number of configurations in which the two rolls sum to a given number, e.g.: $(\# \otimes \#)_{(+)} : 1 \mapsto 0, 2 \mapsto 1$.

Theorem (Law of the Unconscious Statistician)

Formulae for reparameterising integration and measurement:

$$\mu : \Omega, \alpha : X^\Omega, k : X \rightsquigarrow Y \vdash \int d\mu_\alpha k = \int d\mu(k \circ \alpha)$$

$$\mu : \Omega, \alpha : X^\Omega, f : \mathbb{W}^X \vdash \int d\mu_\alpha f = \int d\mu(f \circ \alpha)$$

$$\mu : \Omega, \alpha : X^\Omega, E : \mathcal{B}_X \vdash \text{Ce}_{\mu_\alpha}[E] = \text{Ce}_\mu[\alpha^{-1}[E]] = \text{Ce}_{\omega \sim \mu}[\alpha \omega \in E]$$

Compositional building blocks for modelling

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 $\oint : \mathbf{D}X \times (\mathbf{D}Y)^{\mathbf{D}X} \rightarrow \mathbf{D}Y$

Standard vocabulary: concepts concerning products

Let $\pi_i : \prod_{i \in I} X_i \rightarrow X_i$ be the i -th projection.

Joint distribution: $\mu : \mathbf{D}(X \times Y)$, $\mu : \mathbf{D}(\prod_{i \in I} X_i)$

Marginal distribution: the law of a projection:

$$\mu : \mathbf{D}\left(\prod_{i \in I} X_i\right) \vdash \mu_{\pi_i} : \mathbf{D}X_i$$

Sometimes refers to any law of a r.e..

Marginalisation: the action of calculating a marginal distribution by integrating all other components.

Exercise

$$\mu : \mathbf{P}X, \nu : \mathbf{D}X \vdash (\mu \otimes \nu)_{\pi_2} = \nu$$

Pairing random elements

$$\alpha : X^\Omega, \beta : Y^\Omega \vdash \lambda \omega. (\alpha \omega, \beta \omega) : (X \times Y)^\Omega$$

Independent random elements

The joint law is the product of the marginals:

$$\mu : \mathbf{D}\Omega, \alpha : X^\Omega, \beta : Y^\Omega \vdash \alpha \underset{\mu}{\perp} \beta := \left(\mu_{(\alpha, \beta)} \stackrel{\mathbf{D}(X \times Y)}{=} \mu_\alpha \otimes \mu_\beta \right)$$

More generally, for finite I :

$$\mu : \mathbf{D}\Omega, \alpha_- : (X^\Omega)^I \vdash \underset{\mu}{\perp}_i \alpha_i := \left(\mu_{(\alpha_i)_i} \stackrel{\mathbf{D}(\prod_i X_i)}{=} \bigotimes_{i \in I} \mu_{\alpha_i} \right)$$

Example [Durett]

Model 3 independent coin tosses:

$$\text{Toss} := \{\mathbf{Head}, \mathbf{Tail}\} \quad \Omega := \text{Toss}^3 \quad \mu := \mathbf{U}_{\text{Toss}} \otimes \mathbf{U}_{\text{Toss}} \otimes \mathbf{U}_{\text{Toss}} : \mathbf{P}\Omega$$

The outcome of the i^{th} coin toss is the random element $\pi_i : \Omega \rightarrow \text{Toss}$.

Consider the Boolean proposition in which the i^{th} and j^{th} tosses ($i \neq j$) agree:

$$\text{Same}_{ij} := \lambda\omega. \pi_i\omega = \pi_j\omega : \Omega \rightarrow \mathbb{B}$$

$$\begin{aligned} \text{Calculate: } \Pr_{\mu} [\text{Same}_{12}] &= \overset{\text{LOTUS}}{\downarrow} \Pr_{(x,y) \sim \mu(\pi_1, \pi_2)} [x = y] \\ &= \overset{\text{marginalisation}}{\downarrow} \Pr_{(x,y) \sim \mathbf{U} \otimes \mathbf{U}} [x = y] \\ &= \overset{\text{Fubini}}{\downarrow} \int \mathbf{U}(dx) \Pr_{y \sim \mathbf{U}} [x = y] \\ &= \frac{1}{2} \cdot \Pr_{y \sim \mathbf{U}} [\mathbf{Head} = y] + \frac{1}{2} \cdot \Pr_{y \sim \mathbf{U}} [\mathbf{Tail} = y] = \frac{1}{4} + \frac{1}{4} = \frac{1}{2} \end{aligned}$$

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$$\text{Same}_{ij} := \lambda\omega. \pi_i\omega = \pi_j\omega : \Omega \rightarrow \mathbb{B}$$

Therefore $\mu_{\text{Same}_{12}} = \mathbf{U}_{\mathbb{B}}$ and similarly $\mu_{\text{Same}_{ij}} = \mathbf{U}_{\mathbb{B}}$ for $i \neq j$.

π_1 , Same_{12} , and Same_{13} determine π_2, π_3 , so:

$$\Pr_{\omega \sim \mu} [\text{Same}_{12}\omega = \text{True}, \text{Same}_{13}\omega = \text{True}]$$

Fubini-Tonelli

$$\begin{aligned} &\downarrow \\ &= \int \mathbf{U}_{\text{Toss}}(db_1) \Pr_{(b_2, b_3) \sim (\mathbf{U} \otimes \mathbf{U})} [\text{Same}_{12}(b_1, b_2, b_3) = \text{True}, \text{Same}_{13}(b_1, b_2, b_3) = \text{True}] \\ &= \frac{1}{2} \Pr_{(b_2, b_3) \sim (\mathbf{U} \otimes \mathbf{U})} [\text{Same}_{12}(\text{Head}, b_2, b_3) = \text{True}, \text{Same}_{13}(\text{Head}, b_2, b_3) = \text{True}] \\ &+ \frac{1}{2} \Pr_{(b_2, b_3) \sim (\mathbf{U} \otimes \mathbf{U})} [\text{Same}_{12}(\text{Tail}, b_2, b_3) = \text{True}, \text{Same}_{13}(\text{Tail}, b_2, b_3) = \text{True}] \\ &= \frac{1}{2} \cdot \frac{1}{2} \cdot \frac{1}{2} + \frac{1}{2} \cdot \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{4} \end{aligned}$$

and similarly we get $\frac{1}{4}$ in all other cases.

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The outcome of the i^{th} coin toss is the random element $\pi_i : \Omega \rightarrow \text{Toss}$.

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Therefore $\mu_{\text{Same}_{12}} = \mathbf{U}_{\mathbb{B}}$ and similarly $\mu_{\text{Same}_{ij}} = \mathbf{U}_{\mathbb{B}}$ for $i \neq j$. So:

$$\mu_{(\text{Same}_{12}, \text{Same}_{13})} = \mathbf{U}_{\mathbb{B} \times \mathbb{B}} = \mathbf{U}_{\mathbb{B}} \otimes \mathbf{U}_{\mathbb{B}} = \mu_{\text{Same}_{12}} \otimes \mu_{\text{Same}_{13}}$$

So $\text{Same}_{12} \perp_{\mu} \text{Same}_{13}$ even though their values depend on the outcome of the first toss.

Distribution preservation

Distribution space (Ω, μ)

A type Ω equipped with a distribution $\mu : \mathbf{D}\Omega$. Define **probability space** analogously.

Distribution preserving function

$f : (\Omega_1, \mu_1) \rightarrow (\Omega_2, \mu_2)$ is a function whose is the co domain distribution:

$$f : \Omega_1 \rightarrow \Omega_2 \quad (\mu_1)_f = \mu_2$$

$\mu : \mathbf{D}X$ is **invariant** under $f : X \rightarrow X$ when $f : (X, \mu) \rightarrow (X, \mu)$ is dist. preserving.

Example

Consider the swapping function: $\text{swap} := (\lambda (x, y). (y, x)) : X \times Y \rightarrow Y \times X$. Then, for each $\mu : \mathbf{D}X$, $\nu : \mathbf{D}Y$, swapping is distribution preserving function:

$$\text{swap} : (X \times Y, \mu \otimes \nu) \rightarrow (Y \times X, \nu \otimes \mu)$$

swap is invariant in the case $X = Y$ and $\mu = \nu$.

Density and scaling

Density

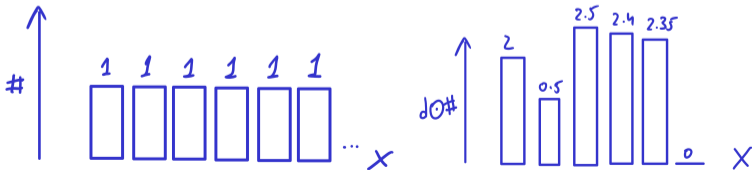
over X is any weight function $f : X \rightarrow \mathbb{W}$.

Density scaling

We can scale a distribution by a density:

$$f : \mathbb{W}^X, \mu : \mathbb{D}X \vdash f \odot \mu := \int \mu(dx)(f, x) \cdot \delta_x : \mathbb{D}X$$

Scaling does not lift to probability distributions: $\|f \odot \mu\| \neq 1$ even if $\|\mu\| = 1$.



Density and scaling

Density

over X is any weight function $f : X \rightarrow \mathbb{W}$.

Density scaling

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Scaling does not lift to probability distributions: $\|f \odot \mu\| \neq 1$ even if $\|\mu\| = 1$.

Warning!

The types of distributions and densities over X in the **discrete** model are close, but still **different**. They coincide on **countable** types, so people often confused them.

Types help us keep them separate.

Density and absolute continuity

Having density

This concept has several names in the literature:

$$\mu, \nu : \mathbf{DX}, f : \mathbb{W}^X \vdash \left(f = \frac{d\mu}{d\nu} \right) := (\mu = f \odot \nu) : \mathbf{Prop}$$

- ▶ f is the **density** of μ w.r.t. ν
- ▶ f is a **Radon-Nikodym derivative** of μ w.r.t. ν .

Absolute continuity

μ is **absolutely continuous** w.r.t. ν when μ has a density w.r.t. ν :

$$\mu, \nu : \mathbf{DX} \vdash (\mu \ll \nu) := \exists f : \mathbb{W}^X. f = \frac{d\mu}{d\nu} : \mathbf{Prop}$$

Example

The **uniform distribution** is absolutely continuous w.r.t. the **counting measure** over the same support. Indeed, it has these two densities:

$$S : \mathcal{P}_{\text{fin}}(X) \vdash \left(\lambda x. \frac{1}{|S|} \right), \left(\lambda x. \begin{cases} x \in S : \frac{1}{|S|} \\ x \notin S : 0 \end{cases} \right) = \frac{d\mathbf{U}_S}{d\#_S}$$

These two densities are different, but they agree on the support, motivating the following concept.

Almost certain event

is one we can assert without changing the distribution:

$$\frac{\Gamma \vdash \mu : DX \quad \Gamma, x : X \vdash M : \mathbb{B}}{\Gamma \vdash \mu(dx) \text{ almost certainly } M := [M] \odot \mu = \mu : \text{Prop}}$$

For probabilities we define:

$$\frac{\Gamma \vdash \mu : PX \quad \Gamma, x : X \vdash M : \mathbb{B}}{\Gamma \vdash \mu(dx) \text{ almost surely } M := (\text{cast } \mu)(dx) \text{ almost certainly } M : \text{Prop}}$$

Existence and almost-sure uniqueness of densities

Theorem (Radon-Nikodym)

For **probability** distributions, we characterise absolute continuity as follows:

$$\mu, \nu : \mathbf{P}X \vdash (\mu \ll \nu) \iff \forall E : \mathcal{B}_X. \Pr_{\nu}[E] = 0 \implies \Pr_{\mu}[E] = 0$$

In that case, if $f, g = \frac{d\mu}{d\nu}$ then $\nu(dx)$ almost surely $f x = g x$.

In the **discrete model**, this characterisation amounts to $\text{supp } \mu \subseteq \text{supp } \nu$.

Example

For all countable X , we have:

$$\forall \mu : \mathbf{D}X. \mu \ll \#_X$$

Indeed, apply the Radon-Nikodym theorem, since $\text{supp } \# = X$.

Constructively, direct calculation shows: $(\lambda x. \mu x) = \frac{d\mu}{d\#}$.

Compositional building blocks for modelling

- ▶ Affine combinations of distributions
- ▶ Product measures (\otimes) : $\mathbf{D}X \times \mathbf{D}Y \rightarrow \mathbf{D}(X \times Y)$
- ▶ Random elements and their laws (push-forward measure):
 $(\lambda(\mu, \alpha) . \mu_\alpha) : \mathbf{D}\Omega \times X^\Omega \rightarrow \mathbf{D}X$

NB:

Standard vocabulary

- ▶ Joint and marginal distributions
- ▶ Independence
- ▶ Distribution/probability preservation and invariance
- ▶ Density and absolute continuity
- ▶ Almost certain/sure properties

- ▶ Dirac kernel $\delta_- : X \rightarrow \mathbf{D}X$
- ▶ Kock integration
 $\oint : \mathbf{D}X \times (\mathbf{D}Y)^{\mathbf{D}X} \rightarrow \mathbf{D}Y$

Part 1: the **discrete** model (now)

- ▶ Motivation
- ▶ Language of probability and distribution
- ▶ Discrete model
- ▶ Simply-typed probability
- ▶ Dependently-typed probability

Part 2: the **full** model

- ▶ Borel sets and measurable spaces
- ▶ Quasi-Borel spaces
- ▶ Type structure & standard Borel spaces
- ▶ Dependently-typed structure
- ▶ Integration



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Example: Binomial kernels

We've defined, for every $n \in \mathbb{N}$, the binomial kernel:

$$\vdash \mathbf{B}_n : \mathbb{1} \rightsquigarrow \mathbf{Fin}(1 + n)$$

We will now look at **dependent-type** structure which allows us to view these as one kernel internally:

$$n : \mathbb{N} \vdash \mathbf{B}_n : \mathbb{1} \rightsquigarrow \mathbf{Fin}(1 + n)$$

Family over an indexing set I

consists of a sequence $X_{\cdot} = (X_i)_{i \in I}$ of sets.

We call each set X_i the **fibre over i** .

Family F

a pair $F = (I, X_{\cdot})$ consisting of (indexing) set I and a family X_{\cdot} over it.

Notation: $F = I \vdash X_{\cdot}$

$$= i : I \vdash X_i.$$

Example

The family $n : \mathbb{N} \vdash \mathbf{Fin} n$ has \mathbb{N} as the indexing set. The fibre over $n \in \mathbb{N}$ is:

$$\mathbf{Fin} n := \{0, 1, \dots, n - 1\}$$

Family model

Family over an indexing set I

consists of a sequence $X_{\cdot} = (X_i)_{i \in I}$ of sets.

We call each set X_i the **fibre over i** .

Family F

a pair $F = (I, X_{\cdot})$ consisting of (indexing) set I and a family X_{\cdot} over it.

Notation: $F = I \vdash X_{\cdot}$

$$= i : I \vdash X_i.$$

Family map

$(\theta, f_{\cdot}) : (I \vdash X_{\cdot}) \rightarrow (J \vdash Y_{\cdot})$ is a pair of a function between the indexing sets and a sequence of functions between the corresponding fibres:

$$\theta : I \rightarrow J \quad (f_i : X_i \rightarrow Y_{\theta i})_{i \in I}$$

Notation: $\theta \vdash f_{\cdot}$. We won't use these maps explicitly, but they are the foundation.

Dependent elements $i : I \vdash M : X_i$

in family $i : I \vdash X_i$ are I -indexed sequences of elements from the corresponding fibres:

$$(M \in X_i)_{i \in I}$$

Example

We have the elements:

$$n : \mathbb{N} \vdash 0, \dots, n - 1 : \mathbf{Fin} \, n$$

Subsumption

Every simple type becomes a family by ignoring the dependency through the constant family, e.g., $i : I \vdash \mathbb{N}$ and $i : I \vdash 42 : \mathbb{N}$.

Simple functions

Fibred exponential

of two families over the same indexing set $i : I \vdash X_i, Y_i$ is the family:

Family of distributions

$$i : I \vdash X_i \rightarrow Y_i$$

over a family $i : I \vdash X_i$ is the family:

$$i : I \vdash \mathbf{D}X_i$$

Its sub-family of fibred **probability** distributions:

$$i : I \vdash \mathbf{P}X_i$$

Both have a **Dirac** distribution:

$$i : I \vdash \delta_- : X_i \rightarrow \mathbf{D}X_i \quad i : I \vdash \delta_- : X_i \rightarrow \mathbf{P}X_i$$

Extension and dependent pairs

Extension

of indexing set I by a **variable** of the family $i : I \vdash X_i$ is the (indexing) set:

$$\coprod_{i \in I} X_i := \bigcup_{i \in I} \{i\} \times X_i = \left\{ (i, x) \in I \times \bigcup_{i \in I} X_i \mid x \in X_i \right\}$$

Notation: $(i : I, x : X_i) := \coprod_{i \in I} X_i$ and we'll often write i, x instead of (i, x) .

Dependent pairs

$$\frac{i : I \vdash X_i \quad i : I, x : X_i \vdash Y_{i,x}}{i : I \vdash (x : X_i) \times (Y_{i,x}) := \coprod_{x \in X_i} Y_{i,x}}$$

Dependent functions

we identify a function f with a tuple $(f x)_x$ as usual:

$$\frac{i : I \vdash X_i \quad i : I, x : X_i \vdash Y_{i,x}}{i : I \vdash ((x : X) \rightarrow Y_{i,x}) := \prod_{x \in X} Y_{i,x}}$$

Dependent kernels $i : I \vdash k : (x : X_i) \rightsquigarrow Y_{i,x}$

are dependent elements:

$$i : I \vdash k : (x : X_i) \rightarrow \mathbf{D}Y_{i,x}$$

Dependent **stochastic** kernels $i : I \vdash k : (x : X_i) \rightsquigarrow Y_{i,x}$ are similarly:

$$i : I \vdash k : (x : X_i) \rightarrow \mathbf{P}Y_{i,x}$$

Dependent Kock integral

$$i : I, \mu : \mathbf{D}X_i, k : (x : X_i) \rightsquigarrow Y_{i,x} \vdash \int d\mu k : \mathbf{D}Y_{i,x}$$

and in the **discrete model** we define it for i, μ, k as in the simply-typed case:

$$\left(\int d\mu k\right)y := \sum_{x \in X_i} \mu x \cdot k(x; y) : \mathbb{W}$$

Through the identification $\mathbb{W} \cong \mathbf{D}\mathbf{1}$ and characteristic functions, we reduce dependent Lebesgue integration and measurement to dependent Kock integration:

$$i : I, \mu : \mathbf{D}X_i, f : (x : X_i) \rightarrow \mathbb{W} \vdash \int d\mu f : \mathbb{W} \qquad i : I, \mu : \mathbf{D}X_i, E : \mathcal{B}_{X_i} \vdash \mathbf{C}e_{\mu}[E] : \mathbb{W}$$

$$\int d\mu f = \sum_{x \in X} \mu x \cdot f x \qquad \mathbf{C}e_{\mu}[E] = \sum_{x \in E} \mu x$$

Random variables

Let $\bar{\mathbb{R}} := [-\infty, \infty]$ be the extended real line.

Signed and unsigned random variable

in a probability space (Ω, μ) are random elements $\alpha : \Omega \rightarrow \bar{\mathbb{R}}$ and $\alpha : \Omega \rightarrow \mathbb{W}$.

The **positive** and **negative parts** are unsigned random variables $\alpha^\pm : \bar{\mathbb{R}}^\Omega \rightarrow \mathbb{W}^\Omega$:

$$\alpha^+ := \lambda\omega. \max(\alpha\omega, 0) = [\alpha \geq 0] \cdot |\alpha| \quad \alpha^- := \lambda\omega. -\min(\alpha\omega, 0) = [\alpha \leq 0] \cdot |\alpha|$$

An unsigned r.v. α is **Lebesgue integrable** when its Lebesgue integral is finite:

$$\int d\mu\alpha < \infty.$$

For a (signed) r.v. α , when either α^+ or α^- is Lebesgue integrable, we define:

$$\mu : \mathbb{D}X, \alpha : \bar{\mathbb{R}}^X, \int d\mu\alpha^+, \int d\mu\alpha^- < \infty \vdash \int d\mu\alpha := \int d\mu\alpha^+ - \int d\mu\alpha^-$$

A signed variable is **Lebesgue integrable** when both its parts are Lebesgue integrable.

Random variable spaces

Lebesgue integrability is a Boolean property:

$$\mu : \mathbf{DX}, \alpha : X \rightarrow \overline{\mathbb{R}} \vdash \alpha \text{ integrable} := \int d\mu \alpha^+ < \infty \wedge \int d\mu \alpha^- < \infty : \mathbb{B}$$

Lebesgue spaces ensemble

is the family:

$$i : I, p : [1, \infty), \mu : \mathbf{PX}_i \vdash \mathcal{L}_p(X_i, \mu) := \{ \alpha : X_i \rightarrow \overline{\mathbb{R}} \mid \alpha^p \text{ integrable} \}$$

Every fibre has a vector space structure and a norm (almost a Banach space!):

$$i : I, p : [1, \infty), \mu : \mathbf{PX}_i, \alpha : \mathcal{L}_p(X_i, \mu) \vdash \|\alpha\|_p := \sqrt[p]{\mathbb{E}_\mu [|\alpha|^p]} : \mathbb{W}$$

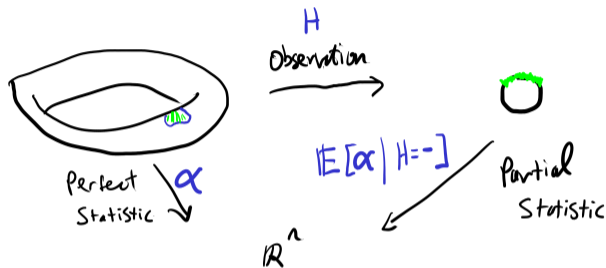
and the fibre **2** has an inner product (almost a Hilbert space!):

$$i : I, \mu : \mathbf{PX}_i, \alpha, \beta : \mathcal{L}_2(X_i, \mu) \vdash (\alpha, \beta) := \sqrt{\mathbb{E}_\mu [\alpha \cdot \beta]} : \mathbb{W}$$

Conditioning à la Kolmogorov

Situation:

- ▶ Statistical model $\mu : \mathcal{D}\Omega$
(voters in the next election)
- ▶ Perfect statistic $\alpha : \Omega \rightarrow \mathbb{R}$
(expected winning candidate)
- ▶ Observation $H : \Omega \rightarrow X$
(poll voting intention)



Conditional expectation of α along H w.r.t μ

Statistic $\beta : X \rightarrow \mathbb{R}$ that 'best' approximates $H \circ \alpha$ statistically. Halmos and Doob's definition: any measurement we make of β agrees with measurement of α :

$$\mu : \mathcal{D}\Omega, H : \Omega \rightarrow X, \alpha : \mathcal{L}_1(\Omega, \mu), \beta : \mathcal{L}_1(X, \mu_H) \vdash$$

$$\left(\beta = \mathbb{E}_{\mu} [\alpha | H = -] \right) := \left(\forall \varphi : \mathcal{L}_1(X, \mu_H). \int d\mu_H \beta \cdot \varphi = \int d\mu \alpha \cdot (\varphi \circ H) \right) \quad : \text{Prop}$$

Theorem (Kolmogorov)

Every random variable has a conditional expectation:

$$\mu : \mathbf{D}\Omega, H : \Omega \rightarrow X, \alpha : \mathcal{L}_1(\Omega, \mu) \vdash \quad \exists \beta : \mathcal{L}_1(X, \mu_H). \beta = \mathbb{E}_{\mu} [\alpha | H = -]$$

Therefore:

Corollary (Internal conditional expectation)

In the **discrete model** we have a dependent function:

$\mathbb{E}_{-} [- | - = -] :$

$$(\mu : \mathbf{D}\Omega) \rightarrow (H : \Omega \rightarrow X) \rightarrow (\alpha : \mathcal{L}_1(\Omega, \mu)) \rightarrow \left\{ \beta : \mathcal{L}_1(X, \mu_H) \mid \beta = \mathbb{E}_{\mu} [\alpha | H = -] \right\}$$

Conditional probability

of event is a conditional expectation of its characteristic function:

$$\mu : \mathbf{P}\Omega, H : \Omega \rightarrow X, E : \mathcal{B}_\Omega, \beta : \mathcal{L}_1(X, \mu_H) \vdash$$
$$\left(\beta = \Pr_\mu [E | H = -] \right) := \left(\beta = \mathbb{E}_{\omega \sim \mu} [\omega \in E | H = -] \right) : \mathbf{Prop}$$

Regular conditional probability

a kernel that agrees with the conditional expectation of the characteristic functions:

$$\mu : \mathbf{P}\Omega, H : \Omega \rightarrow X, k : X \rightsquigarrow \Omega \vdash$$
$$\left(k = \Pr_\mu [- | H = -] \right) := \left(\forall E \in \mathcal{B}_\Omega. k(-; E) = \mathbb{E}_{\omega \sim \mu} [\omega \in E | H = -] \right) : \mathbf{Prop}$$

Conditioning via disintegration

Kolmogorov's theorem does **not** ensure the existence of a regular conditional probability, although the constructive, discrete, definition does.

Disintegration Problem (warning: conflicting terminologies in literature)

Input: probability distribution $\mu : \mathbb{P}\Omega$, measurable map $H : \Omega \rightarrow \Theta$
induce law $\nu := \mu_H : \mathbb{P}\Theta$

Output: probability kernel $k : \Theta \rightsquigarrow \Omega$ such that: $\mu = \int d\nu k$.

We call k a **disintegration** of μ along H .

Proposition

Consider a probability kernel $k : \Theta \rightsquigarrow \Omega$. TFAE:

- ▶ k is a disintegration of μ along $H : \Omega \rightarrow \Theta$;
- ▶ k is a regular conditional probability kernel of μ conditioned on H .

Conditioning via disintegration

Fibred disintegration of $\mu : \mathbb{P} \left(\coprod_{\Theta} \Omega \right)$ (non-standard terminology and formulation)

a partial dependent kernel $k : (\theta : \Theta) \rightsquigarrow \Omega_{\perp}$, defined μ_{dep} -a.s., that disintegrates μ along the first projection $\text{dep} : \left(\coprod_{\Theta} \Omega \right) \rightarrow \Theta$:

$$\mu : \mathbb{P} \left(\coprod_{\Theta} \Omega \right), k : \Theta \rightsquigarrow \Omega_{\perp} \vdash k \text{ disintegrates fibres of } \mu :=$$
$$\mu_{\text{dep}}(\text{Dom}(k)) = 1, \mu = \int \text{d}\mu_{\text{dep}} k : \text{Prop}$$

In the **discrete model** we have an internal disintegration:

$$-\dagger : \left(\mu : \mathbb{P} \left(\coprod_{\Theta} \Omega \right) \right) \rightarrow \{k : (\theta : \Theta) \rightsquigarrow \Omega_{\perp} \mid k \text{ disintegrates } \mu \text{ along } \text{dep}\}$$
$$\text{Dom} \left(\mu^{\dagger} \right) := \{ \theta \mid \mu_{\text{dep}} \theta > 0 \} \quad \mu^{\dagger} := \lambda \theta. \frac{1}{\mu_{\text{dep}} \theta} \odot \mu \upharpoonright_{\text{dep}^{-1}[\theta]}$$

Bayes's Theorem (adapted from Williams)

Let:

- ▶ $\lambda : \mathbb{P}(X \times \Theta)$ be a joint probability distribution.
- ▶ $\mu : \mathbb{D}X$, $\nu : \mathbb{D}\Theta$ be distributions such that $\lambda \ll \mu \otimes \nu$ $X \xleftarrow{\alpha := \pi_1} X \times \Theta \xrightarrow{H := \pi_2} \Theta$
- ▶ $w_{\alpha, H} = \frac{d\lambda}{d\mu \otimes \nu} : X \times \Theta \rightarrow \mathbb{W}$ a Radon-Nikodym derivative

Observation 1

- ▶ $w_\alpha := \lambda x. \int \nu(d\theta) w_{\alpha, H}(x, \theta) : X \rightarrow \mathbb{W}$ then: $w_\alpha = \frac{d\lambda_\alpha}{d\mu}$
- ▶ $w_H := \lambda \theta. \int \mu(dx) w_{\alpha, H}(x, \theta) : \Theta \rightarrow \mathbb{W}$ then: $w_H = \frac{d\lambda_H}{d\nu}$

Observation 2

Let: $w_\alpha(- | H = -) : X \times \Theta \rightarrow \mathbb{W}$ $w_\alpha(x | H = \theta) := \begin{cases} w_H \theta > 0 : & \frac{w_{\alpha, H}(x, \theta)}{w_H \theta} \\ \text{otherwise:} & 0 \end{cases}$

$\lambda_{\alpha | H = -} : \Theta \rightsquigarrow X$ $\lambda_{\alpha | H = \theta} := \lambda_\alpha(- | H = \theta) \odot \nu$. Then:

$$\lambda_{\alpha | H = -} = \Pr_\lambda[- | H = -] \quad (\text{Bayes's formula})$$

Part 1: the **discrete** model

- ▶ Motivation
- ▶ Language of probability and distribution
- ▶ Discrete model
- ▶ Simply-typed probability
- ▶ Dependently-typed probability

Part 2: the **full** model (now)

- ▶ Borel sets and measurable spaces
- ▶ Quasi-Borel spaces
- ▶ Type structure & standard Borel spaces
- ▶ Dependently-typed structure
- ▶ Integration



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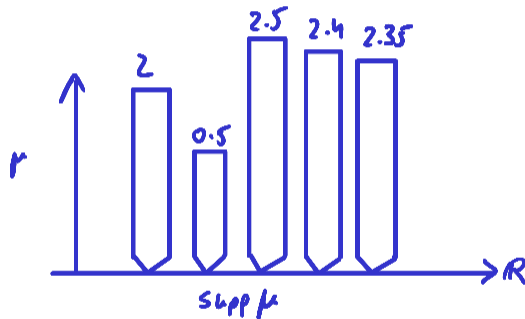
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From histograms to measures

The **discrete** model expresses **histograms** only.

Also want **continuous** distributions:

- ▶ lengths
- ▶ areas
- ▶ volumes



Theorem (Vitali 1905)

There is no reasonable generalisation of 'length' that measures all subsets of the real line—there is no function $\lambda : \mathcal{P}\mathbb{R} \rightarrow \mathbb{W}$ satisfying:

$$\lambda[a, b] = (b - a)$$

(generalise length)

$$\lambda(s + [E]) = \lambda E$$

(translation invariance)

$$\lambda(\bigsqcup_{i=0}^{\infty} E_n) = \sum_{i=0}^{\infty} \lambda E_n$$

(σ -additivity)

Takeaway

$\mathcal{B}_{\mathbb{R}} := \mathcal{P}\mathbb{R}$ as in the **discrete** model excludes **length, area, volume** as distributions.

\implies need a different model

Workaround

Only measure **well-behaved** subsets:

Borel subsets $\mathcal{B}_{\mathbb{R}} \subseteq \mathcal{P}\mathbb{R}$

smallest **σ -field** containing all **open intervals**:

$\overline{\emptyset} \in \mathcal{B}_{\mathbb{R}}$	$\frac{E \in \mathcal{B}_{\mathbb{R}}}{E^c \in \mathcal{B}_{\mathbb{R}}}$	$\frac{E_n \in \mathcal{B}_{\mathbb{R}}}{\bigcup_{n \in \mathbb{N}} E_n \in \mathcal{B}_{\mathbb{R}}}$	$\frac{a, b \in \mathbb{R}}{(a, b) \in \mathcal{B}_{\mathbb{R}}}$
(empty set)	(complements)	(countable unions)	(intervals)

Examples

- ▶ Countable discrete subsets are Borel:

$$\{r\} = \bigcap_{\varepsilon \in \mathbb{Q}_{>0}} (r - \varepsilon, r + \varepsilon) \in \mathcal{B}_{\mathbb{R}} \quad , \quad I \text{ countable} \implies I = \bigcup_{i \in I} \{i\}$$

- ▶ Any interval is Borel, e.g.: $[a, b) = (a, b) \cup \{a\}$

Measurable space $M = (\underline{M}, \mathcal{B}_M)$

set of **points** $a \in \underline{M}$ equipped with a **σ -field** $\mathcal{B}_M \subseteq \mathcal{P}\underline{M}$:

$$\overline{\emptyset} \in \mathcal{B}_\mathbb{R}$$

(empty set)

$$\frac{E \in \mathcal{B}_\mathbb{R}}{E^c \in \mathcal{B}_\mathbb{R}}$$

(complements)

$$\frac{E_n \in \mathcal{B}_\mathbb{R}^{\mathbb{N}}}{\bigcup_{n \in \mathbb{N}} E_n \in \mathcal{B}_\mathbb{R}}$$

(countable unions)

Examples

▶ Discrete spaces: $\overline{I}^{\text{Meas}} := (I, \mathcal{P}I)$

▶ Sub-spaces: $\frac{S \subseteq \underline{M}}{S_M := (S, [\mathcal{B}_M] \cap S)}$ i.e., $\mathcal{B}_{S_M} := \{E \cap S \mid E \in \mathcal{B}_M\}$, e.g., $[0, \infty) \hookrightarrow \mathbb{R}$

▶ Products: $\mathcal{B}_{\prod_{i \in I} M_i} := \sigma \bigcup_{i \in I} \pi_i^{-1} [\mathcal{B}_{M_i}] = \sigma \left\{ \times_{i \in I} E_i \mid \begin{array}{l} E_- \in \prod_{i \in I} \mathcal{B}_{M_i}, \\ \exists J \subseteq_{\text{countable}} I. \\ \forall j \notin J. E_j = \underline{M}_j \end{array} \right\}$, e.g.: \mathbb{R}^n

Borel measurable function $f : M \rightarrow K$

function sending points to points and measurable subsets to measurable subsets:

$$f : \underline{M} \rightarrow \underline{K} \quad \mathcal{B}_M \ni f^{-1}[E] \iff E \in \mathcal{B}_K$$

Examples

▶ $(+), (\cdot) : \mathbb{R}^2 \rightarrow \mathbb{R}$

▶ $|\cdot|, \sin : \mathbb{R} \rightarrow \mathbb{R}$

▶ any continuous function $\mathbb{R}^n \rightarrow \mathbb{R}$

▶ any function out of a discrete space: $\frac{f : I \rightarrow M}{f : \bar{I} \rightarrow M}$

Category Meas

Objects M : measurable spaces

Arrows $f : M \rightarrow K$: Borel measurable functions

$$\frac{}{\text{id} := (\lambda x.x) : M \rightarrow M} \qquad \frac{f : M \rightarrow K \quad g : K \rightarrow L}{g \circ f : (\lambda x.g(f x)) : M \rightarrow L}$$

Categorical structure

Products, coproducts/disjoint unions, subspaces, projective and injective limits / categorical limits and colimits are all fine.

Theorem (Aumann'61)

There are no measurable spaces of Borel subsets nor of measurable functions over \mathbb{R} . In detail, there are no σ -fields $\mathcal{B}_{\mathcal{B}_{\mathbb{R}}}$ and $\mathcal{B}_{\mathbb{R} \rightarrow \mathbb{R}}$ such that, letting $\mathcal{B}_{\mathbb{R}}$ and $\mathbb{R} \rightarrow \mathbb{R}$ be the corresponding measurable spaces, the following functions are measurable:

- ▶ Membership testing:

$$(\in) := \left(\lambda r.E. \begin{cases} r \in E : & \mathbf{True} \\ \text{otherwise:} & \mathbf{False} \end{cases} \right) : \mathbb{R} \times \mathcal{B}_{\mathbb{R}} \rightarrow \overline{\{\mathbf{True}, \mathbf{False}\}}$$

- ▶ Evaluation: $\text{eval} := (\lambda (f, r). f r) : (\mathbb{R} \rightarrow \mathbb{R}) \times \mathbb{R} \rightarrow \mathbb{R}$.

As a consequence, \mathbf{Meas} is not Cartesian closed.

Aumann's Theorem: proof preliminaries

Recall the **Borel hierarchy** over a family of subsets $\mathcal{U} \subseteq \mathcal{P}X$, defined by transfinite induction on $\omega_1 + 1$, the successor of the first uncountable ordinal:

$$\Sigma_\alpha^{\mathcal{U}}, \Pi_\alpha^{\mathcal{U}}, \Delta_\alpha^{\mathcal{U}} \subseteq \mathcal{P}X \quad (\alpha \in \omega_1)$$

$$\Sigma_1^{\mathcal{U}} := \mathcal{U}$$

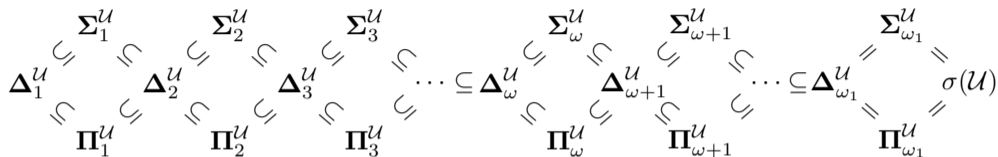
$$\Sigma_{\alpha+1}^{\mathcal{U}} := \left\{ \bigcup_{i \in I} A_i \mid I \subseteq \mathbb{N}, A_i \in \mathcal{U} \cup \bigcup_{\beta \leq \alpha} \Pi_\beta^{\mathcal{U}} \right\} \quad (1 \leq \alpha \in \omega_1)$$

$$\Sigma_\gamma^{\mathcal{U}} := \bigcup_{\beta < \gamma} \Sigma_\beta^{\mathcal{U}} \quad (1 \leq \gamma \text{ a limit ordinal in } \omega_1)$$

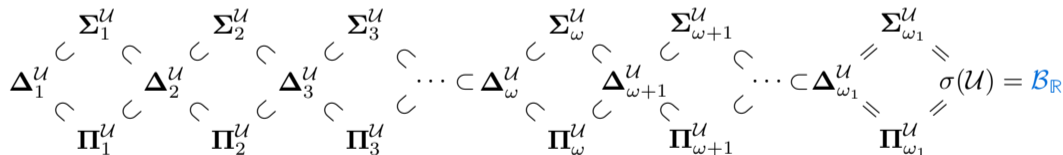
$$\Pi_\alpha^{\mathcal{U}} := [\Sigma_\alpha^{\mathcal{U}}]^c := \{A^c \mid A \in \Sigma_\alpha^{\mathcal{U}}\} \quad \Delta_\alpha^{\mathcal{U}} := \Sigma_\alpha^{\mathcal{U}} \cap \Delta_\alpha^{\mathcal{U}}$$

Aumann's Theorem: proof preliminaries

The Borel hierarchy looks like this in general:



For $\mathcal{U} := \{(a,b) \mid a, b \in \mathbb{R}\}$, the hierarchy does not stabilise before ω_1 :



Rank of $E \in \sigma\mathcal{U}$

first step in which it appears: $\text{Rank}_E := \min \{ \alpha < \omega_1 \mid A \in \Delta_\alpha^U \}$.

Proof

Assume to the contrary there was some σ -field providing a measurable space of Borel subsets $\mathcal{B}_{\mathbb{R}}$ such that membership testing is measurable:

$$(\epsilon) : \mathbb{R} \times \mathcal{B}_{\mathbb{R}} \rightarrow \overline{\{\mathbf{True}, \mathbf{False}\}} \quad \text{NB: } \mathcal{B}_{\mathbb{R} \times \mathcal{B}_{\mathbb{R}}} = \sigma([\mathcal{B}_{\mathbb{R}}] \times [\mathcal{B}_{\mathcal{B}_{\mathbb{R}}}]$$

Let $\alpha := \text{Rank}(\epsilon)^{-1}[\mathbf{True}] < \omega_1$, and find $E \in \mathcal{B}_{\mathbb{R}}$ with $\text{Rank}_E > \alpha$. Then:

$$\begin{aligned} \alpha < \text{Rank } E &= \text{Rank} \left(((\epsilon) \circ (-, E))^{-1}[\mathbf{True}] \right) = \text{Rank} \left((-, E)^{-1} \left((\epsilon)^{-1}[\mathbf{True}] \right) \right) \\ &\leq \text{Rank} \left((\epsilon)^{-1}[\mathbf{True}] \right) = \alpha \end{aligned}$$

So $\alpha < \alpha$, a contradiction, and the postulated σ -field cannot exist. A similar proof replacing E with its characteristic function proves eval cannot be measurable. ■

Some higher-order structure in Meas

Sequences

By generalities, $(\bar{I} \rightarrow M) = \prod_{i \in I} M$. For countable I , we use $\bar{I} \rightarrow M$ for sequences.

Example

A sequence $a_- : \mathbb{N} \rightarrow \mathbb{R}$ is **Cauchy** when its tail elements tend infinitesimally close:

$$\forall \varepsilon > 0. \exists N \in \mathbb{N}. \forall m, n > N. |a_n - a_m| < \varepsilon$$

The Cauchy property characterises convergence to a finite limit. We can define the Cauchy property through quantification over countable sets:

$$\mathbf{Cauchy} \in \mathcal{B}_{\mathbb{N} \rightarrow \mathbb{R}} \quad \mathbf{Cauchy} := \bigcap_{\varepsilon \in \mathbb{Q}_{>0}} \bigcup_{N \in \mathbb{N}} \bigcap_{m, n \in \mathbb{N}} \{a_- \in \underline{\mathbb{N} \rightarrow \mathbb{R}} \mid |a_n - a_m| < \varepsilon\}$$

$$\text{measurability through type-checking: } = \left\{ a_- \in \underline{\mathbb{N} \rightarrow \mathbb{R}} \mid \forall \varepsilon : \mathbb{Q}_{>0}. \exists N : \mathbb{N}. \forall m, n : \mathbb{N}. m, n > N \implies |a_n - a_m| < \varepsilon \right\}$$

Measurability through type-checking

With a few simple building blocks:

$$\limsup : (\mathbb{N} \rightarrow \mathbb{R}) \rightarrow [-\infty, \infty]$$

$$\limsup^{-1}[b, \infty] = \{a \in \underline{\mathbb{N} \rightarrow \mathbb{R}} \mid \forall n : \mathbb{N}. \exists m : \mathbb{N}. m > n \wedge a_m \geq b\} \in \mathcal{B}_{\mathbb{N} \rightarrow \mathbb{R}}$$

we can discharge measurability through type-checking:

$$\lim : \mathbf{Cauchy} \rightarrow \mathbb{R} \quad \lim a_- := \limsup a_-$$

$$\mathbf{Vanishing} := \left\{ r_- : \mathbb{N} \rightarrow \mathbb{R}_{>0} \mid \lim_{n \rightarrow \infty} r_n = 0 \right\} \in \mathcal{B}_{\mathbb{N} \rightarrow \mathbb{R}}$$

$$\mathbf{approx}_- : \mathbf{Vanishing} \times \mathbb{R} \rightarrow (\mathbb{N} \rightarrow \mathbb{Q}) \quad \text{such that: } |r - \mathbf{approx}_{\Delta_-} r n| < \Delta_n$$

However, not all operations of interest support this technique:

$$\times \quad \limsup : (\mathbb{N} \rightarrow \mathbb{R} \rightarrow [-\infty, \infty]) \rightarrow (\mathbb{R} \rightarrow [-\infty, \infty]) \quad \limsup f_- := \lambda x. \limsup_{n \rightarrow \infty} f_n x$$

as they are intrinsically higher-order.

Goal

Measurability by type-checking! Want to extend the model without sacrificing/compromising the language of probability we developed.

Challenge: compositionality

For **higher-order** building blocks, classical measure theory **defers** measurability proofs until we resume 1st-order fragment.

Some probabilistic concepts are inherently higher-order, and classical measure theory makes them 2nd-class.

Sbs \leftrightarrow Meas

A measurable space $S \in \mathbf{Meas}$ is **standard Borel** when there is a measurable isomorphism $S \cong E$ for some $E \in \mathcal{B}_{\mathbb{R}}$.

Concrete spaces are standard

- ▶ Discrete countable spaces \bar{I}
- ▶ Countable products of standard spaces are standard: $\mathbb{R}^n, \mathbb{N} \rightarrow \mathbb{R}, \mathbb{N} \rightarrow \mathbb{B}, \mathbb{N} \rightarrow \mathbb{N}$.
- ▶ Borel subspaces of standard spaces are standard: $[0,1], \mathbb{R}_{>0}, \mathbb{R}_{\geq 0}$.
- ▶ Countable coproducts of standard spaces are standard: $[0,\infty], [-\infty,\infty]$.

Plan: use a different conservative extension

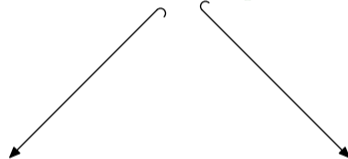
Concrete spaces:

standard Borel spaces

Abstract spaces:

measurable spaces

quasi-Borel spaces



Part 1: the **discrete** model

- ▶ Motivation
- ▶ Language of probability and distribution
- ▶ Discrete model
- ▶ Simply-typed probability
- ▶ Dependently-typed probability

Part 2: the **full** model

- ▶ Borel sets and measurable spaces
- ▶ Quasi-Borel spaces
- ▶ Type structure & standard Borel spaces
- ▶ Dependently-typed structure
- ▶ Integration

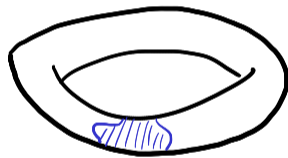


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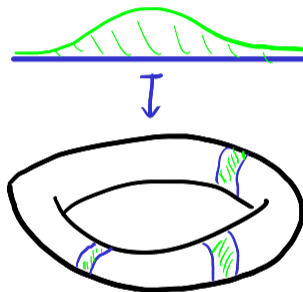


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measurable spaces



quasi Borel spaces



primitive notions: points & **measurable subsets**

points & **random elements**

derived notions: **random elements** $\alpha : \mathbb{R} \rightarrow M$

measurable subsets $E \in \mathcal{B}_X$

Metaphorology¹ \mathcal{R}

over set \underline{X} of **points**: subset of functions $\alpha : \mathbb{R} \rightarrow \underline{X}$ closed under:

constants:

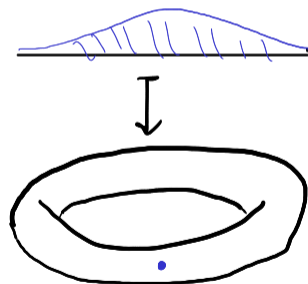
$$\frac{x \in \underline{X}}{\underline{x} := (\lambda r. x) \in \mathcal{R}}$$

precomposition:

$$\frac{\alpha \in \mathcal{R} \quad \varphi \in \text{Meas}(\mathbb{R}, \mathbb{R})}{(\alpha \circ \varphi) \in \mathcal{R}}$$

recombination:

$$\frac{I \subseteq \mathbb{N} \quad E_- : I \rightarrow \mathcal{B}_{\mathbb{R}} \quad \mathbb{R} = \bigsqcup_{i \in I} E_i \quad \alpha_- : I \rightarrow \mathcal{R}}{[E_i. \alpha_i]_{i \in I} := (\lambda r \in E_i. \alpha_i r) \in \mathcal{R}}$$



¹ $\mu\epsilon\tau\alpha$ ('meta', across) and $\varphi\epsilon\rho\omega$ ('phero', to carry).

Core definitions

Metaphorology¹ \mathcal{R}

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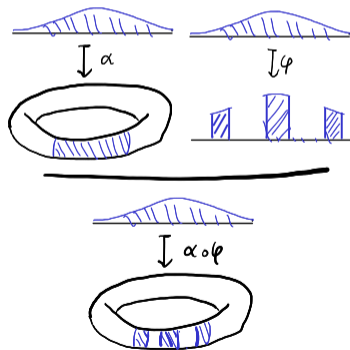
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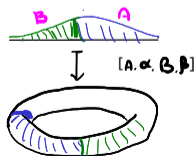
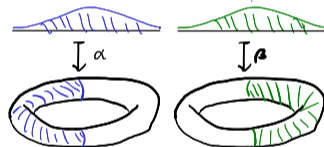
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Quasi Borel space (qbs) X
set of points \underline{X} equipped with
a metaphorology \mathcal{R}_X over it

recombination:

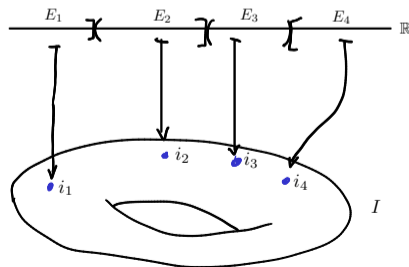
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Examples

- ▶ **real line** qbs: $\mathbb{R} = (\mathbb{R}, \mathbf{Meas}(\mathbb{R}, \mathbb{R}))$
- ▶ **underlying qbs** of a measurable space: $M = (\underline{M}, \mathbf{Meas}(\mathbb{R}, M))$
- ▶ **indiscrete** qbs over a set: Indiscrete $I = (I, \mathbb{R} \rightarrow I)$
- ▶ **discrete** qbs over a set: $\bar{I} = (I, \mathcal{R}_{\bar{I}})$ with the σ -**simple** metaphorology, consisting of recombinations of constants:

$$\mathcal{R}_{\bar{I}} := \left\{ [E_j \cdot i_j]_{j \in J} \mid J \subseteq \mathbb{N}, E_- : J \rightarrow \mathcal{B}_{\mathbb{R}}, \mathbb{R} = \biguplus_{j \in J} E_j, i_- : J \rightarrow I \right\}$$



Core definitions

Example validation: qbs underlying \mathbb{W}

elements: $\underline{\mathbb{W}} = [0, \infty]$; random elements: $\mathcal{R}_{\mathbb{W}} = \mathbf{Meas}(\mathbb{R}, \mathbb{W})$.

$$\text{constants: } \underline{w}^{-1}[E] = \begin{cases} w \in E : \mathbb{R} \in \mathcal{B}_{\mathbb{R}} \\ w \notin E : \emptyset \in \mathcal{B}_{\mathbb{R}} \end{cases} \implies \underline{w} \in \mathbf{Meas}(\mathbb{R}, \mathbb{W}).$$

$$\text{precomposition: } \frac{\alpha \in \mathbf{Meas}(\mathbb{R}, \mathbb{W}) \quad \varphi \in \mathbf{Meas}(\mathbb{R}, \mathbb{R})}{(\varphi \circ \alpha) \in \mathbf{Meas}(\mathbb{R}, \mathbb{W})}$$

since composition in \mathbf{Meas} is function composition

$$\text{recombination: } \frac{I \subseteq \mathbb{N} \quad E_- : I \rightarrow \mathcal{B}_{\mathbb{R}} \quad \alpha_- : I \rightarrow \mathbf{Meas}(\mathbb{R}, \mathbb{W}) \quad \mathbb{R} = \bigsqcup_{i \in I} E_i}{\forall F \in \mathcal{B}_{\mathbb{W}} : [E_i \cdot \alpha_i]_{i \in I}^{-1}[F] = \bigcup_{i \in I} \alpha_i^{-1}[F] \cap E_i \in \mathcal{B}_{\mathbb{R}}}$$

Indeed:

$$r \in \text{LHS} \iff [E_i \cdot \alpha_i]_{i \in I} r \in F \iff \exists i \in I. r \in E_i \wedge \alpha_i r \in F \iff r \in \text{RHS}$$

Example validation: discrete qbs over I

elements: I ; random elements: σ -simple functions.

constants: $\underline{i} = [\mathbb{R}.\underline{i}]_{j=1}^1 \in \mathcal{R}_{\bar{I}}$

precomposition: $[E_i.\underline{i}_j]_j \circ \varphi = [\varphi^{-1}[E_i].\underline{i}_j]_{j, \varphi^{-1}[E_j] \neq \emptyset} \in \mathcal{R}_{\bar{I}}$

recombination: slightly more fiddly, but similar:

$$\left[F_\ell \cdot [E_{\ell,j}.\underline{i}_{\ell,j}]_{j \in J_\ell} \right]_{\ell \in L} = \left[F_\ell \cap E_{\ell,j}.\underline{i}_{\ell,j} \right]_{\substack{(\ell,j) \in \coprod_{\ell \in L} J_\ell \\ F_\ell \cap E_{\ell,j} \neq \emptyset}} \in \mathcal{R}_{\bar{I}}$$

The category of quasi Borel spaces

Quasi-Borel measurable function $f : X \rightarrow Y$

function sending points to points and random elements to random elements:

$$f : \underline{X} \rightarrow \underline{Y} \quad \mathcal{R}_X \ni \alpha \implies f \circ \alpha \in \mathcal{R}_Y$$

Examples

- ▶ **constant** functions $\underline{b} : X \rightarrow Y$, as send any r.e. to a σ -simple function.
- ▶ **σ -simple** functions $\alpha : \mathbb{R} \rightarrow X$, as send any r.e. to a σ -simple function.
- ▶ **Borel measurable** functions are quasi-Borel measurable, by precomposition:

$$f \in \mathbf{Meas}(M, K) \quad \implies \quad f : M \rightarrow K$$

Category Qbs

consists of quasi Borel spaces and quasi-Borel measurable functions, with identity functions and function composition.

Language of **probability** & **distribution** (full model)

X type of **values/outcomes**: quasi Borel space

DX type of **distributions/measures** over X :

$PX \subseteq DX$ sub-type of **probability distributions** over X :

$\mathcal{B}_X \subseteq \mathcal{P}X$ type of **events**—subsets we wish to measure:

\mathbb{W} type of **weights**: values in $[0, \infty]$

\int, \mathbb{E} Lebesgue integration and the expectation operation

Type judgements describe well-formed values/outcomes of a given type, e.g.:

$$\mu : DX, E : \mathcal{B}_X \vdash \underset{\mu}{\text{Ce}}[E] : \mathbb{W}$$

are measurable functions

Propositions describe properties of well-formed values/outcomes of a given type, e.g.:

$$y_1, y_2 : Y \vdash y_1 \stackrel{Y}{=} y_2 : \text{Prop} \quad \mu : PX, E : \mathcal{B}_X \vdash \underset{\mu}{\text{cast Pr}}[E] = \underset{\mu}{\text{Ce}}[E]$$

are arbitrary set-theoretic propositions

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Proposition

Take I set and X qbs:

- ▶ every function $f : I \rightarrow \underline{X}$ is measurable as $f : \bar{I} \rightarrow X$.
- ▶ every function $\underline{X} \rightarrow I$ is measurable as $f : X \rightarrow \text{Indiscrete } I$.

Proposition

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- ▶ every function $\underline{X} \rightarrow I$ is measurable as $f : X \rightarrow \text{Indiscrete } I$.

Proof

Every function $f : I \rightarrow \underline{X}$ sends σ -simple functions to σ -simple functions:

$$f \circ [E_j \cdot \underline{i}_j]_{j \in J} = [E_j \cdot \underline{f i}_j]_{j \in J} \in \mathcal{R}_X$$

The function $f \circ \alpha$ is always a random element in Indiscrete I . ■

Expectation management

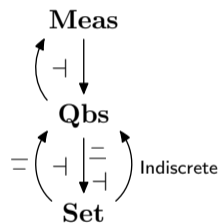
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Adjoint situation

- ▶ limits and colimits are as in **Set**
- ▶ slogan: every measurable space is carried by a qbs.



Simple type structure

Product $\prod_{i \in I} X_i \xrightarrow{\pi_j} X_j$

correlated random elements metaphorology:

$$\mathcal{R}_{\prod_{i \in I} X_i} := \left\{ (\alpha_i)_{i \in I} : \mathbb{R} \rightarrow \prod_{i \in I} X_i \mid \forall i : I. \alpha_i \in \mathcal{R}_{X_i} \right\}$$

Coproduct $\coprod_{i \in I} X_i \xleftarrow{\iota_j} X_j$

metaphorology generated from component metaphorologies:

$$\begin{aligned} \mathcal{R}_{\coprod_{i \in I} X_i} &:= \mathcal{R} \left(\bigcup_{i \in I} (\iota_i \circ) [\mathcal{R}_{X_i}] \right) \\ &= \left\{ [E_j \cdot \alpha_j]_{j \in J} : \mathbb{R} \rightarrow \prod_{i \in I} X_i \mid \begin{array}{l} J \subseteq \mathbb{N}, i_- : J \rightarrow I, E_- : J \rightarrow \mathcal{B}_{\mathbb{R}}, \\ \mathbb{R} = \biguplus_{j \in J} E_j, \forall j. \alpha_j \in \mathcal{R}_{X_{i_j}} \end{array} \right\} \end{aligned}$$

Simple type structure

Function space $(X \rightarrow Y) \times X \xrightarrow{\text{eval}} X$

elements: measurable functions; random elements: curried measurable functions;
evaluation as in **Set**:

$$\begin{aligned} \underline{X \rightarrow Y} &:= \mathbf{Qbs}(X, Y) & \mathcal{R}_{X \rightarrow Y} &:= \text{curry}_X [\mathbf{Qbs}(\mathbb{R} \times X, Y)] \\ \text{eval}(f, a) &:= f a & &= \left\{ \varphi : \mathbb{R} \rightarrow \mathbf{Qbs}(X, Y) \mid \text{uncurry } \varphi := (\lambda (r, a) . (\varphi r) a) \right. \\ & & & \left. : \mathbb{R} \times X \rightarrow Y \right\} \end{aligned}$$

Simple type structure

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$$\underline{X \rightarrow Y} := \mathbf{Qbs}(X, Y) \quad \mathcal{R}_{X \rightarrow Y} := \text{curry}_X [\mathbf{Qbs}(\mathbb{R} \times X, Y)]$$

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Meas vs. Qbs

Using the adjunction (to be established), we have:

$$\begin{array}{ccc}
 (L(\mathbb{R} \rightarrow \mathbb{R})) \times \mathbb{R} & \xrightarrow{\text{Aumann's Thm}} & \mathbb{R} \\
 \uparrow (L\pi_1, L\pi_2) = (\lambda(f, a). (f, a)) & \text{--- } \times \text{---} & \parallel \\
 L((\mathbb{R} \rightarrow \mathbb{R}) \times \mathbb{R}) & \xrightarrow{\text{Leval}} & L\mathbb{R}
 \end{array}$$

$$\begin{array}{c}
 \mathbf{Meas} \\
 \uparrow L \quad \downarrow U \\
 \mathbf{Qbs}
 \end{array}$$

These allow us to interpret a **simple type theory**. We use a **shallow embedding**².

Simple contexts

$$\blacktriangleright x_1 : X_1, \dots, x_n : X_n := \prod_{i=1}^n X_i$$

$$\blacktriangleright (\Gamma \vdash M : X) := \Gamma \xrightarrow{M} X$$

Measurability by type-checking

Use simple type structure as well as primitive space constructions to define measurable **building blocks**. Instead of proving that a function is measurable, build it from smaller measurable functions using the shallow embedding.

²Scibior et al.[POPL'2018] use a deep embedding, e.g.

Proposition

For $X \in \mathbf{Qbs}$:

$$\underline{\mathbb{R} \rightarrow X} = \mathcal{R}_X$$

Proof

(\supseteq). Take $\alpha \in \mathcal{R}_X$. By precomposition, $\alpha : \mathbb{R} \rightarrow X$. So $\alpha \in \underline{\mathbb{R} \rightarrow X}$.

(\subseteq). Take $\alpha \in \underline{\mathbb{R} \rightarrow X}$. Since $\text{id} \in \mathbf{Qbs}(\mathbb{R}, \mathbb{R}) = \mathcal{R}_{\mathbb{R}}$, $\alpha = \alpha \circ \text{id} \in \mathcal{R}_X$. ■

Random element space \mathcal{R}_X

the function space $\mathbb{R} \rightarrow X$.

Proposition

For a measurable function $e : X \rightarrow Y$, TFAE, and we then write $e : X \hookrightarrow Y$:

- ▶ e is a **subspace embedding**: a random element $\beta \in \mathcal{R}_Y$ that lifts pointwise along e extends a unique random element along e :

$$(\forall r : \mathbb{R}. \exists a \in \underline{X}. e a = \beta r) \implies \exists ! \alpha \in \mathcal{R}_X. e \circ \alpha = \beta$$

- ▶ e is injective, and Y 's metaphorology determines X 's:

$$\mathcal{R}_X = \{ \alpha : \mathbb{R} \rightarrow \underline{X} \mid e \circ \alpha \in \mathcal{R}_Y \}$$

- ▶ e is a **strong monomorphism**, i.e., right-orthogonal to all epis.

The subspace embedding $\underline{\mathbf{True}} : \mathbb{1} \hookrightarrow \text{Indiscrete} \{ \mathbf{True}, \mathbf{False} \} =: \mathbf{Prop}$ is a **strong subobject classifier**.

Subspaces

We internalise any set-theoretic property $\varphi : \underline{X} \rightarrow \{\mathbf{True}, \mathbf{False}\}$ into a measurable function $x : X \vdash \varphi : \mathbf{Prop}$, forming the subspace over $\{a \in \underline{X} \mid \varphi a = \mathbf{True}\}$:

$$\frac{x : X \vdash \varphi : \mathbf{Prop}}{\text{coerce}_\varphi : \{x : X \mid \varphi\} \hookrightarrow X} \qquad \frac{\Gamma \vdash M : X \quad \Gamma \vdash \varphi [x \mapsto M]}{\Gamma \vdash \text{lift}_\varphi M : \{x : X \mid \varphi\}}$$

Examples

- ▶ $\mathbb{N} \hookrightarrow \mathbb{Q} \hookrightarrow \mathbb{R}_{\geq 0} \hookrightarrow \mathbb{R}$
- ▶ $\mathbb{W} \hookrightarrow [-\infty, \infty]$
- ▶ Iverson bracket $[-] : \mathbb{B} := \overline{\{\mathbf{True}, \mathbf{False}\}} \hookrightarrow \mathbb{R}$
- ▶ $\mathbb{B} \not\hookrightarrow \mathbf{Prop}$.

Proposition

Let $e : X \hookrightarrow Y$ be a subspace embedding. TFAE, and we say e is **Borel** and write $e : X \vDash Y$:

- ▶ for every $\alpha \in \mathcal{R}_Y$, $\alpha^{-1}[e[X]] \in \mathcal{B}_R$;
- ▶ the characteristic function is measurable: $[- \in e[X]] : Y \rightarrow \mathbb{B}$.

The Borel subspace embedding **True** : $\mathbb{1} \vDash \mathbb{B}$ is a **Borel subspace embedding classifier**.

Space of Borel subsets \mathcal{B}_X

the function space $X \rightarrow \mathbb{B}$. It is an internal σ -field:

$$\emptyset : \mathcal{B}_X \quad -^c : \mathcal{B}_X \rightarrow \mathcal{B}_X \quad \bigcup, \bigcap : (I \rightarrow \mathcal{B}_X) \rightarrow \mathcal{B}_X \quad (I \subseteq \mathbb{N})$$

precisely because \mathbb{B} is an internal σ -algebra in **Sbs**.

Examples

- ▶ Borel subsets of \mathbb{R} as a quasi Borel space are the usual Borel subsets
- ▶ As we've seen in **Meas**, the Cauchy sequences **Cauchy** $\leftrightarrow (\mathbb{N} \rightarrow \mathbb{R})$.
- ▶ For discrete spaces, $\underline{\mathcal{B}}_I = \mathcal{P}I$.
- ▶ For indiscrete spaces $\underline{\mathcal{B}}_{\text{Indiscrete } I} = \{\emptyset, I\}$
- ▶ The left adjoint to **Meas** \rightarrow **Qbs** is given by $X \mapsto (\underline{X}, \mathcal{B}_X)$.
- ▶ For measurable space M , the subsets $\mathcal{B}_{\mathcal{B}_M}$ are the **Borel-on-Borel** subsets from **classical descriptive set theory** [Sabok et al. '21].

Non-Examples [Sabok et al. '21]

The following subspace embeddings are not Borel:

$$\{E \in \mathcal{B}_{\mathbb{R}} \mid E \neq \emptyset\}, \{E \in \mathcal{B}_{\mathbb{R}} \mid E \text{ is open}\} \hookrightarrow \mathcal{B}_{\mathbb{R}} \quad \{(E, F) \in \mathcal{B}_{\mathbb{R}}^2 \mid E \subseteq F\} \hookrightarrow \mathcal{B}_{\mathbb{R}}^2$$

Partial map $f : X \multimap Y$

measurable function $X \rightarrow Y \amalg \{\perp\} =: Y_{\perp}$.

We internalise the partial map space as $(X \multimap Y) := (X \rightarrow Y_{\perp})$.

The **domain of definition** of such a partial map is a Borel subset $\text{Dom}(f) \varepsilon X$.

It internalises $\text{Dom}(-) : (X \multimap Y) \rightarrow \mathcal{B}_X$.

We use internal partial maps to define the space of distributions.

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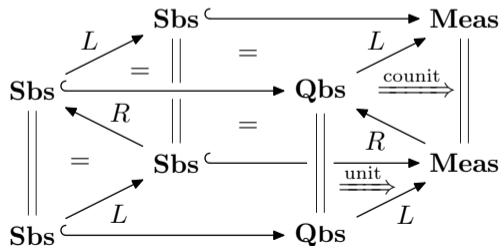
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$\mathbf{Sbs} \leftrightarrow \mathbf{Qbs}$

A quasi Borel space space $S \in \mathbf{Qbs}$ is **standard Borel** when there is a measurable isomorphism $S \cong E$ for some $E \in \mathcal{B}_{\mathbb{R}}$.

Proposition

The adjunction between \mathbf{Meas} and \mathbf{Qbs} restricts to an isomorphism between the corresponding categories of standard Borel spaces:



Standard Borel spaces have good properties.
E.g.: they have disintegrations along arbitrary maps.

Theorem (well-known)

A measurable space is standard iff its σ -field is generated by a **Polish topology**:
completely metrizable separable topology.

Quasi Borel spaces offer an alternative methodology for showing a space is standard:

classical theory

quasi Borel spaces

- ▶ Equip the space with a topology or a σ -field.
- ▶ Show it is standard, typically by metrizing it.
- ▶ Show compatibility with operations of interest, e.g., evaluation is measurable.
- ▶ Use a suitable space, e.g., for evaluation.
- ▶ Exhibit an isomorphism to a standard space.
- ▶ Optionally, study a compatible metric.

Example (well-known)

Let $\mathbf{C}_0[a,b]$ be the quasi-Borel space of continuous functions over $[a,b]$:

$$\mathbf{C}_0[a,b] := \{f : [a,b] \rightarrow \mathbb{R} \mid f \text{ continuous}\} \hookrightarrow ([a,b] \rightarrow \mathbb{R})$$

Immediately, without need of further proof, we have a quasi-Borel space that supports measurable evaluation, after coercion, and measurable abstraction, subject to lifting.

Example (well-known)

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Representation as standard space

A continuous function is determined uniquely by its values on rationals, a countably-long vector of real numbers:

$$-|_{\mathbb{Q} \cap [a,b]} : \mathbf{C}_0 \rightarrow (\mathbb{Q} \cap [a,b] \rightarrow \mathbb{R})$$

Example (well-known)

Let $\mathbf{C}_0[a,b]$ be the quasi-Borel space of continuous functions over $[a,b]$:

$$\mathbf{C}_0[a,b] := \{f : [a,b] \rightarrow \mathbb{R} \mid f \text{ continuous}\} \hookrightarrow ([a,b] \rightarrow \mathbb{R})$$

Representation as standard space

The key is to characterise the image as a Borel subset:

Theorem (Cantor)

A function $f : [a,b] \rightarrow \mathbb{R}$ is continuous iff it is uniformly continuous.

So lift $-|$ to:

$$\begin{aligned} -|_{\mathbb{Q} \cap [a,b]} &:= (\lambda f. \text{lift } \lambda q. f q) : \mathbf{C}_0 \rightarrow \left\{ r_- : \mathbb{Q} \cap [a,b] \rightarrow \mathbb{R} \mid \forall \varepsilon : \mathbb{Q}_{>0}. \exists \delta : \mathbb{Q}_{>0}. \forall p, q : \mathbb{Q} \cap [a,b]. \right. \\ &\quad \left. |p - q| \leq \delta \implies |r_p - r_q| \leq \varepsilon \right\} \\ &=: D \hookrightarrow (\mathbb{Q} \cap [a,b] \rightarrow \mathbb{R}) \in \mathbf{Sbs} \end{aligned}$$

Example (well-known)

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Representation as standard space

Conversely, define the inverse by taking the limit³:

$$\llbracket - \rrbracket := \lambda r . \lambda x . \lim_{n \rightarrow \infty} r \text{ approx } \lambda m . \frac{1}{1+m} x n : D \rightarrow \mathbf{C}_0$$

We now have a standard Borel space equipped with operations of interests, such as abstraction and evaluation.

³I'm glossing over the edge case $x = b$, but a case-split will deal with it.

To fully reproduce the classical account, exhibiting a separable, complete metric that generates the σ -field, we typically need to do more or less the same calculations. To that end, the following two concepts break the task down.

Compatible metric

over a qbs X is a metric that is also measurable $d : X \times X \rightarrow \mathbb{W}$. It has **measurable limits** when the limit function is a measurable partial function $\text{lim} : (\mathbb{N} \rightarrow X) \rightarrow X_{\perp}$.

Theorem

Let d be a compatible metric with measurable limits over a qbs X .

If d is separable, then $\mathcal{B}_d = \mathcal{B}_X$.

Example

The **uniform convergence metric** $d(f, g) := \sup_{x \in [a, b]} |f x - g x|$ is compatible over $\mathbf{C}_0[a, b]$, since we can define it equivalently as $d(f, g) := \sup_{x \in \mathbb{Q} \cap [a, b]} |f x - g x|$. It has measurable limits, since it is complete and its limits are taken pointwise. It is separable by Weierstrass's approximation theorem.

Part 1: the **discrete** model

- ▶ Motivation
- ▶ Language of probability and distribution
- ▶ Discrete model
- ▶ Simply-typed probability
- ▶ Dependently-typed probability

Part 2: the **full** model

- ▶ Borel sets and measurable spaces
- ▶ Quasi-Borel spaces
- ▶ Type structure & standard Borel spaces
- ▶ **Dependently-typed structure**
- ▶ Integration



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Core advantages

A **convenient** setting for dependently-typed probability:

- ▶ possesses enough type constructions.
- ▶ strictly preserves substitution.
- ▶ smooth internalisation and externalisation.

Main ideas

- ▶ Use equivalent structure to locally Cartesian-closed structure, guaranteeing enough type constructions.
- ▶ Use families, deferring most issues involving type dependency to meta-level, and focus on measurability structure and requirements.

Core definitions

Γ -fibred metaphorology \mathcal{R}^- over $\gamma : \underline{\Gamma} \vdash \underline{X}_\gamma$

\mathcal{R}_Γ -indexed family of subsets $v : \mathcal{R}_\Gamma \vdash \mathcal{R}^v \subseteq (r : \mathbb{R}) \rightarrow \underline{X}_{vr}$ closed under:

fibred constants:

$$\frac{\gamma \in \underline{\Gamma} \quad a \in \underline{X}_\gamma}{\underline{a} := (\lambda r. a) \in \mathcal{R}^\gamma}$$

fibred precomposition:

$$\frac{\alpha \in \mathcal{R}^v \quad \varphi \in \mathbf{Qbs}(\mathbb{R}, \mathbb{R})}{(\alpha \circ \varphi) \in \mathcal{R}^{v \circ \varphi}}$$

fibred recombination:

$$\frac{I \subseteq \mathbb{N} \quad E_- : I \rightarrow \mathcal{B}_\mathbb{R} \quad \mathbb{R} = \bigsqcup_{i \in I} E_i \quad \alpha_- : (i : I) \rightarrow \mathcal{R}^{v_i}}{[E_i \cdot \alpha_i]_{i \in I} \in \mathcal{R}^{[E_i \cdot v_i]_{i \in I}}}$$

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Quasi Borel family (qbf) $\Gamma \vdash X_-$

a family $\gamma : \underline{\Gamma} \vdash \underline{X}_\gamma$ equipped with a Γ -fibred metaphorology $\mathcal{R}_{X_-}^-$ over \underline{X}_- .

Examples

- ▶ Every qbs X is a qbf $\Gamma \vdash X$ with $(\underline{X})_\gamma := X$ and $\mathcal{R}_X^v := \mathcal{R}_X$.
- ▶ We will later see that intervals are qbfs $a, b : [-\infty, \infty] \vdash [a, b], [a, b), (a, b), (a, b)$ with $\mathcal{R}^{a \mapsto \alpha, b \mapsto \beta}$ those measurable functions $\varphi \in \mathbf{Qbs}(\mathbb{R}, [-\infty, \infty])$ such that, for all $r \in \mathbb{R}$, φr is in the interval delimited by αr and βr .
- ▶ For measurable map $d \downarrow_{\Gamma}^X$, the **preimage** qbf:

$$\gamma : \Gamma \vdash d^{-1}[\gamma] \quad \mathcal{R}_{d^{-1}[-]}^v := \left(\begin{array}{c} \mathcal{R}_X \\ d \downarrow \\ \mathcal{R}_\Gamma \end{array} \right)^{-1} [v] \subseteq \mathcal{R}_X$$

Core definitions

Qbf map $(\theta \vdash f_-) : (\gamma : \Gamma \vdash X_\gamma) \rightarrow (\delta : \Delta \vdash Y_\delta)$

a family map $(\theta \vdash f_-) : (\gamma : \underline{\Gamma} \vdash \underline{X}_\gamma) \rightarrow (\delta : \underline{\Delta} \vdash \underline{Y}_\delta)$ such that:

$$\mathcal{R}_X^v \ni \alpha \quad \Longrightarrow \quad (f_- \overset{\theta}{\circ} \alpha) := (\lambda r. f_{\theta r}(\alpha r)) \in \mathcal{R}_Y^{\theta \circ v}$$

A **vertical map** $\Gamma \vdash f_- : X_- \rightarrow Y_-$ is a qbf map of the form:

$$(\text{id}_\Gamma \vdash f_-) : (\Gamma \vdash X_-) \rightarrow (\Gamma \vdash Y_-)$$

Core definitions

Qbf map $(\theta \vdash f_-) : (\gamma : \Gamma \vdash X_\gamma) \rightarrow (\delta : \Delta \vdash Y_\delta)$

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We use only vertical maps explicitly. The totality maps forms the foundation though:

Theorem

The category of qbfs and their maps, **Qbf**, equipped with the functor $(\Gamma \vdash X_-) \mapsto \Gamma : \mathbf{Qbf} \rightarrow \mathbf{Qbs}$ form a split fibred category.

The Grothendieck construction and the preimage form a fibred adjoint equivalence between this fibred category and the codomain fibration over **Qbs**.

Contextual structure

Substitution $\Gamma \vdash - [\theta]$

measurable function $\theta : \Gamma \rightarrow \Delta$ operates on qbfs and their vertical maps:

$$\frac{\delta : \Delta \vdash X_-}{\gamma : \Gamma \vdash X_- [\theta]_\gamma} \quad \underline{X_- [\theta]_\gamma} := X_{\theta\gamma} \quad \mathcal{R}_{X_- [\theta]}^v := \mathcal{R}_{X_-}^{\theta \circ v}$$

$$\frac{\delta : \Delta \vdash f_\delta : X_\delta \rightarrow Y_\delta}{\gamma : \Gamma \vdash f_- [\theta]_\gamma := f_{\theta\gamma} : X_{\theta\gamma} \rightarrow Y_{\theta\gamma}}$$

Fibred terminal qbf $\Gamma \vdash \mathbb{1}$

given by the terminal qbf considered as a family. NB: $\mathbb{1} [\theta] = \mathbb{1}$.

Terms $\gamma : \Gamma \vdash M_\gamma : X_\gamma$

vertical maps $\Gamma \vdash M_- : \mathbb{1} \rightarrow X_-$. NB: $\frac{\delta : \Delta \vdash M_\delta : X_\delta}{\Gamma \vdash M_- [\theta]_\gamma = M_{\theta\gamma} : X_{\theta\gamma}}$

Context extension $\gamma : \Gamma, x : X_\gamma$

given by the Grothendieck construction:

$$\underline{\gamma : \Gamma, x : X_\gamma} := \{ \gamma, (x \mapsto a) \mid \gamma \in \underline{\Gamma}, a \in \underline{X}_a \} = \coprod_{\gamma \in \underline{\Gamma}} \underline{X}_\gamma$$

$$\mathcal{R}_{\gamma : \Gamma, x : X_\gamma} := \{ (v, \alpha) \mid v \in \mathcal{R}_\Gamma, \alpha \in \mathcal{R}_{X^\nu} \}$$

and we have weakening substitution and variable term:

$$(\gamma : \Gamma, x : X_\gamma) \xrightarrow{\text{weaken}^{x:X}} \Gamma \quad \gamma : \Gamma, x : X_\gamma \vdash x : X_\gamma = X_- [\text{weaken}]_{\gamma, x \mapsto x}$$

Extensional propositional equality $\frac{\gamma : \Gamma \vdash M_\gamma, N_\gamma : X_\gamma}{\gamma : \Gamma \vdash M_\gamma ::= K_\gamma}$

fibred sub-terminal, non-empty only in fibres where $M_\gamma = K_\gamma$:

$$\underline{M ::= K}_\gamma := \begin{cases} M_\gamma() = K_\gamma() : \mathbb{1} \\ \text{otherwise} : \emptyset \end{cases} \quad \mathcal{R}_{M ::= K}^v := \begin{cases} \forall r : \mathbb{R}. M_{vr}() = K_{vr}() : \left\{ \underline{()}\right\} \\ \text{otherwise} : \emptyset \end{cases}$$

$$\frac{\gamma : \Gamma \vdash M_\gamma : X_\gamma}{\gamma : \Gamma \vdash (\text{refl } M)_\gamma := () : (M_- ::= M_-)_\gamma}$$

Contextual structure

Extensional propositional fording

For, e.g., transport, generalise prop. equality to package under equality assumption:

$$\frac{\gamma : \Gamma \vdash M_\gamma, K_\gamma : X_\gamma \quad \gamma : \Gamma, x : X_\gamma \vdash Y_{\gamma, x \mapsto x}}{\gamma : \Gamma \vdash ((M =: x := K) \times Y_{-, x \mapsto x})_\gamma}$$

$$\frac{((M =: x := K) \times Y_{-, x \mapsto x})_\gamma}{\gamma} := \begin{cases} M_\gamma() =: a := K_\gamma() : Y_{\gamma, x \mapsto a} \\ \text{otherwise :} & \emptyset \end{cases}$$

$$\mathcal{R}_{(M=:x:=K) \times Y_{-, x \mapsto x}}^v := \begin{cases} M_v() =: \alpha := K_v() : \mathcal{R}_{Y_-}^{v, x \mapsto \alpha} \\ \text{otherwise :} & \emptyset \end{cases}$$

$$\frac{\Gamma \vdash M : X_- \quad \Gamma \vdash N : Y_{-, x \mapsto M}}{\Gamma \vdash \text{refl } M, N := N_\gamma() : (M =: x := M) \times Y_{-, x \mapsto x}}$$

$$\frac{\Gamma, x : X_- \vdash Y_- \quad \Gamma, y : X_- \vdash Z_- \quad \Gamma \vdash M, K : X_- \quad \Gamma \vdash f_- : Y_{-, x \mapsto M} \rightarrow Z_{-, y \mapsto M}}{\Gamma \vdash \text{match - with } \{f_-\} := (\lambda b. f_{\gamma, x \mapsto M_\gamma} b) : (M =: x := K) \times Y_- \rightarrow Z_{-, y \mapsto K}}$$

Dependent pairs $\gamma : \Gamma \vdash (x : X_\gamma) \times Y_{\gamma, x \mapsto x}$

given by dependent pair family, with correlated fibred random elements:

$$\begin{aligned} \underline{(x : X_-) \times Y_{-, x \mapsto x}}_\gamma &:= (x : \underline{X}_\gamma) \times \underline{Y}_{\gamma, x \mapsto x} \\ \mathcal{R}_{(x : X_-) \times Y_{-, x \mapsto x}}^v &:= \left\{ (\alpha, \beta) \mid \alpha \in \mathcal{R}_{\underline{X}_-}^v, \beta \in \mathcal{R}_{\underline{Y}_-}^{v, x \mapsto \alpha} \right\} \end{aligned}$$

$$\frac{\gamma : \Gamma \vdash M_\gamma : X_\gamma \quad \gamma : \Gamma \vdash K_\gamma : Y_{\gamma, x \mapsto M}}{\gamma : \Gamma \vdash M_\gamma, K_\gamma := (\lambda(). M_\gamma(), K_\gamma()) : (x : X_\gamma) \times Y_{\gamma, x \mapsto x}}$$

$$\frac{\gamma : \Gamma \vdash M : (x : X_\gamma) \times Y_{\gamma, x \mapsto x} \quad \gamma : \Gamma, p : (x : X_\gamma) \times Y_{\gamma, x \mapsto x} \vdash Z_{\gamma, p \mapsto p} \quad \gamma, a : X_\gamma, b : Y_{\gamma, x \mapsto a} \vdash K : Z_{\gamma, p \mapsto a, b}}{\text{match } M \text{ with } \{a, b \Rightarrow K\} := (\lambda(a', b'). K_{\gamma, a \mapsto a', b \mapsto b'}()) : Z_{\gamma, p \mapsto M}}$$

Dependent functions $\gamma : \Gamma \vdash (x : X_\gamma) \rightarrow Y_{\gamma, x \mapsto x}$

given by subfamily of dependent functions that preserve fibred random elements:

$$\begin{aligned} \underline{(x : X_-) \rightarrow Y_{-, x \mapsto x}_\gamma} &:= \left\{ f : (x : X_\gamma) \rightarrow Y_{\gamma, x \mapsto x} \mid \forall \alpha \in \mathcal{R}_{X_-}^\gamma. f \circ \alpha \in \mathcal{R}_{Y_-}^{\gamma, x \mapsto \alpha} \right\} \\ \mathcal{R}_{(x : X_-) \rightarrow Y_{-, x \mapsto x}}^v &:= \left\{ \varphi : (r : \mathbb{R}) \rightarrow \underline{(x : X_-) \rightarrow Y_{-, x \mapsto x}}_{v r} \mid \begin{array}{l} \forall \rho \in \mathbf{Qbs}(\mathbb{R}, \mathbb{R}). \forall \alpha \in \mathcal{R}_{X_-}^{v \circ \rho}. \\ (\lambda r. ((\varphi \circ \rho) r)(\alpha r)) \\ \in \mathcal{R}_{Y_- v \circ \rho, x \mapsto \alpha} \end{array} \right\} \end{aligned}$$

$$\frac{\gamma : \Gamma, x : X \vdash M : Y_{\gamma, x \mapsto x}}{\gamma : \Gamma \vdash \lambda x : X_\gamma. M := (\lambda(). \lambda a. M_{\gamma, x \mapsto a}()) : (x : X_\gamma) \rightarrow Y_{\gamma, x \mapsto x}}$$

$$\frac{\gamma : \Gamma \vdash M : (x : X_\gamma) \rightarrow Y_{\gamma, x \mapsto x} \quad \gamma : \Gamma \vdash K : X_\gamma}{\gamma : \Gamma \vdash M K := (\lambda(). M_\gamma() K_\gamma()) : Y_{\gamma, x \mapsto M}}$$

Theorem ({In/Ext}ternalisation)

Let I be a set and \bar{I} its discrete qbs.

A qbf $\gamma : \Gamma, i : \bar{I} \vdash X_{\gamma, i \mapsto i}$ amounts to an I -indexed family of qbfs $(\gamma : \Gamma \vdash X_{\gamma, i \mapsto j})_{j \in I}$.

Similarly, a term $\gamma : \Gamma, i : \bar{I} \vdash M_{\gamma, i \mapsto i} : X_{\gamma, i \mapsto i}$ amounts to an I -indexed family of terms $(\gamma : \Gamma \vdash M_{\Gamma, i \mapsto j} : X_{\gamma, i \mapsto j})_{j \in I}$.

We can use this theorem to show that fibred I -ary products and coproducts are given by dependent functions and dependent pairs indexed by the discrete qbs \bar{I} :

$$\Gamma \vdash \prod_{i \in I} X_i := (i : \bar{I}) \times X_i, \quad \coprod_{i \in I} X_i := (i : \bar{I}) \rightarrow X_i$$

Fibred exponentials $\gamma : \Gamma \vdash X_\gamma \rightarrow Y_\gamma$

$$\underline{X_- \rightarrow Y_-}_\gamma := \mathbf{Qbs}(X_\gamma, Y_\gamma)$$

$$\mathcal{R}_{X_- \rightarrow Y_-}^v := \left\{ \varphi : (r : \mathbb{R}) \rightarrow \mathbf{Qbs}(X_{vr}, Y_{vr}) \mid \begin{array}{l} \forall \rho \in \mathbf{Qbs}(\mathbb{R}, \mathbb{R}). \forall \alpha \in \mathcal{R}_{X_-}^{v \circ \rho}. \\ (\lambda r. ((\varphi \circ \rho) r)(\alpha r)) \in \mathcal{R}_{Y_-}^{v \circ \rho} \end{array} \right\}$$

By treating the qbsees `Prop` and `B` as qbfs, we can classify fibred subspace embeddings, and Borel subspace embeddings, as in the simply-typed setting. With the fibred coproduct, we can define fibred partial maps, as in the simply-typed setting.

Part 1: the **discrete** model

- ▶ Motivation
- ▶ Language of probability and distribution
- ▶ Discrete model
- ▶ Simply-typed probability
- ▶ Dependently-typed probability

Part 2: the **full** model

- ▶ Borel sets and measurable spaces
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- ▶ Type structure & standard Borel spaces
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We still need:

- ▶ the spaces of distributions $\mathbf{D}X$;
- ▶ their Dirac kernels $\delta_- : X \rightarrow \mathbf{D}X$; and
- ▶ the Kock integrals $\mu : \mathbf{D}X, f : X \rightarrow \mathbf{D}Y \vdash \oint d\mu f : \mathbf{D}Y$

satisfying the appropriate properties.

The rest follows internally, as we have observed in Part I.

Before presenting the ‘right’ choice for $\mathbf{D}X$, we give a few alternatives and explain their short-comings.

Spaces of measures

When we introduced the language of probability, we mentioned three properties of distributions:

$$\mu : DX \vdash \text{Ce}_{\mu}[\emptyset] = 0 \quad (\text{strictness})$$

$$\mu : DX, E, F : \mathcal{B}_X \vdash \text{Ce}_{\mu}[E] = \text{Ce}_{\mu}[E \cap F] + \text{Ce}_{\mu}[E \cap F^c] \quad (\text{disjoint additivity})$$

$$\mu : DX, E_{-} : (\mathcal{B}_X, \subseteq)^{\omega} \vdash \text{Ce}_{\mu}\left[\bigcup_n E_n\right] = \sup_n \text{Ce}_{\mu}[E_n] \quad (\text{Scott continuity})$$

In measure theory, we take a measure to be exactly a function with this properties:

A measure μ over a measurable space M

consists of a (set-theoretic) function $\mu : \mathcal{B}_M \rightarrow \mathbb{W}$ that is strict, disjointly additive, and Scott continuous.

We write $\underline{G}'M$ for the set of measures over M . For a qbs X , define

$$\underline{G}'X := \underline{G}'(X, \mathcal{B}_X).$$

Lebesgue measure $\lambda \in \underline{G}\mathbb{R}$

the unique measure assigning each interval its length: $\lambda[a, b] = b - a$, for all $a \leq b$.

The proof is a little technical, first extending λ to finite unions of intervals, and appealing to the celebrated **Carathéodory's extension theorem**.

See Williams [91], e.g., for more details.

We can equip $\underline{G'X}$ with a metaphorology:

✘ Kernel metaphorology $\mathcal{R}_{G'X}$

the random elements are exactly the kernels from \mathbb{R} to X :

$$\mathcal{R}_{G'X} := \{k : \mathbb{R} \rightarrow \underline{GX} \mid \forall E \in \mathcal{B}_X. (\lambda r. k(r, E) \in \mathbf{Qbs}(\mathbb{R}, \mathbb{W}))\}$$

Were we to choose this metaphorology, we could define Lebesgue $(\mu, f \mapsto \int d\mu f)$, for example, but it will only be measurable in μ , and not in f . This issue is invisible in measure theory as one cannot define measurability in f .

✓ Giry⁴ qbs \mathcal{R}_{GX}

the subspace $\underline{GX} \hookrightarrow (\mathcal{B}_X \rightarrow \mathbb{W})$.

The Giry qbs may have fewer elements $\underline{GR} \subset \underline{G'R}$ and random elements: $\mathcal{R}_{GR} \subset \mathcal{R}_{G'R}$.

⁴Named after Michèle Giry's 1982 measure-theoretic monad for probability distributions.

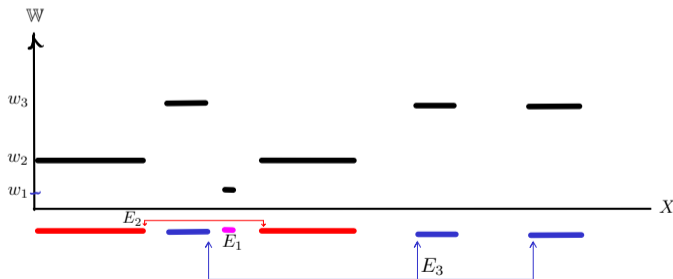
Lebesgue integration

We now proceed to define the Lebesgue integral for the qbs Giry monad.

Simple functions $\varphi : X \rightarrow \mathbb{W}$

finite recombinations of constants:

$$\varphi = [E_i \cdot \underline{w}_i]_{i=1}^n \quad n \in \mathbb{N} \quad E_- : \mathcal{B}_X^n \quad \underline{X} = \bigcup_{i=1}^n E_i \quad \underline{w}_- : \mathbb{W}^n$$



Lebesgue integration

We encode simple functions by recording their Borel partition and their value in each subset of this partition:

The spaces of simple functions and their codes

codes describe the components in the recombination:

$$\mathbf{SimpleFun}_X := \{f : X \rightarrow \mathbb{W} \mid f \text{ simple}\} \quad \mathbf{SimpleCode}_X := \prod_{n \in \mathbb{N}} \mathcal{B}_X^n \times \mathbb{W}^n$$

We arbitrarily chose not to require the subsets to form a partition.

The interpretation of a simple code as a simple function is a measurable function:

$$\llbracket - \rrbracket : \mathbf{SimpleCode}_X \rightarrow \mathbf{SimpleFun} \quad \llbracket (n, E_-, w_-) \rrbracket := \text{lift } \lambda a. \sum_{i=1}^n [a \in E_i] \cdot w_i$$

It is measurable by type-checking.

Lemma

A function $f : \underline{X} \rightarrow \underline{W}$ is measurable iff $f = \lim_{n \rightarrow \infty} f_n$ for some pointwise ω -chain $f_- \in \mathbf{SimpleFun}^\omega$.

Moreover, there is a (dependent) measurable function assigning codes for such approximating sequence:

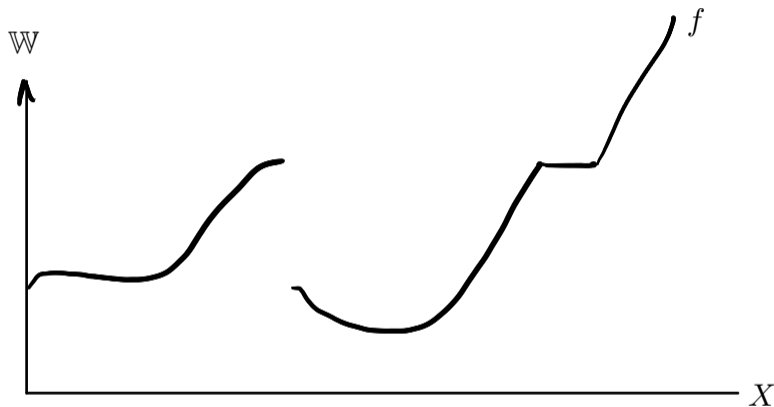
approxSimple_x

$$: (f : X \rightarrow W) \rightarrow (\Delta_- : \mathbf{Vanishing}) \rightarrow (W_- : W^\omega) \rightarrow \lim_{n \rightarrow \infty} W_n ::= \infty \rightarrow$$
$$\left\{ f_- : \mathbb{N} \rightarrow \mathbf{SimpleCode} \left| \begin{array}{l} \forall n. \llbracket f_n \rrbracket \leq \llbracket f_{n+1} \rrbracket \leq f, \forall n, x. \llbracket f_n \rrbracket x \leq W_n, \\ \forall n, x. fx \leq W_n \implies |fx - \llbracket f_n \rrbracket| < \Delta_n \end{array} \right. \right\}$$

- ▶ Δ_- is the monotonically decreasing and vanishing rate of convergence; and
- ▶ W_- is the range of convergence.

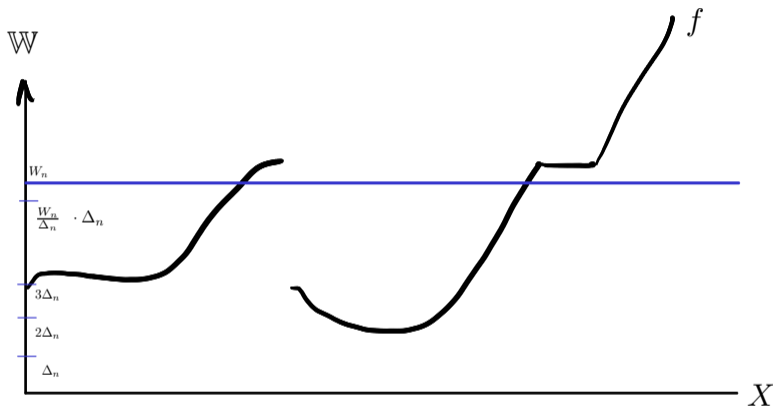
Lebesgue integration

Why?



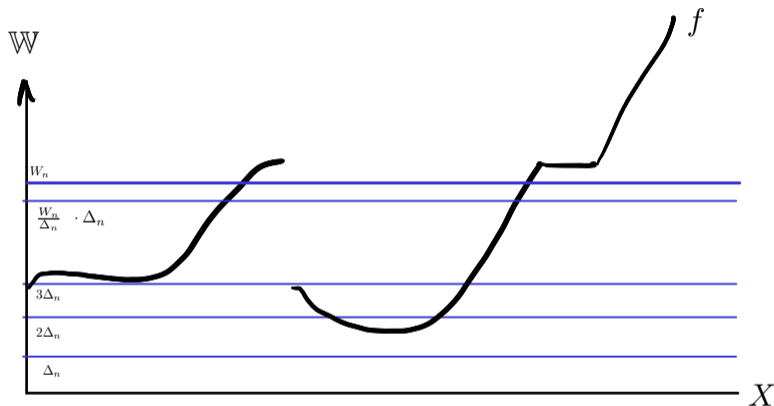
Lebesgue integration

Why?



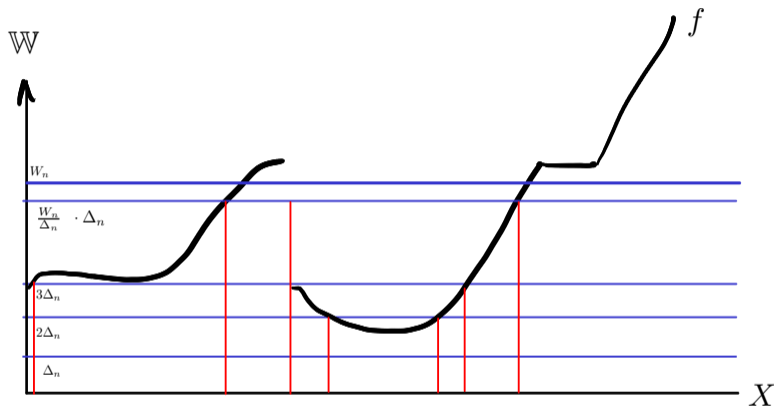
Lebesgue integration

Why?



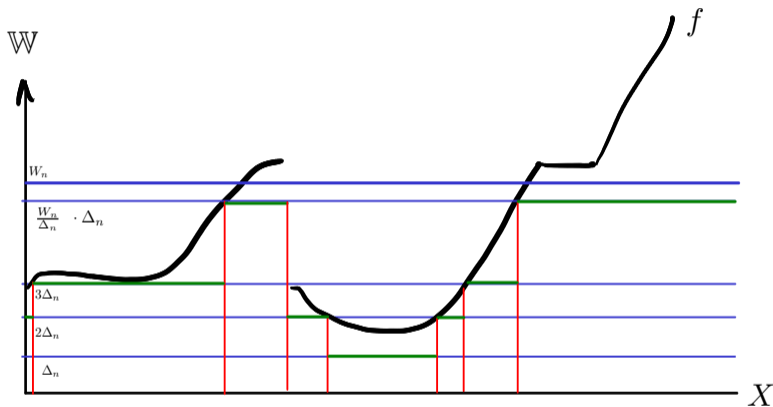
Lebesgue integration

Why?



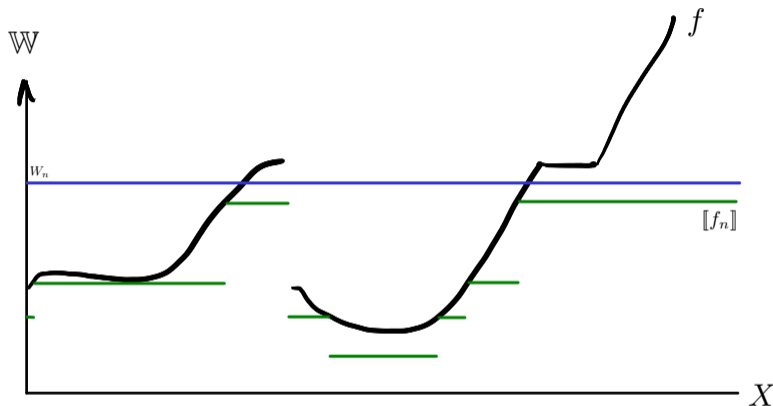
Lebesgue integration

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Lebesgue integration

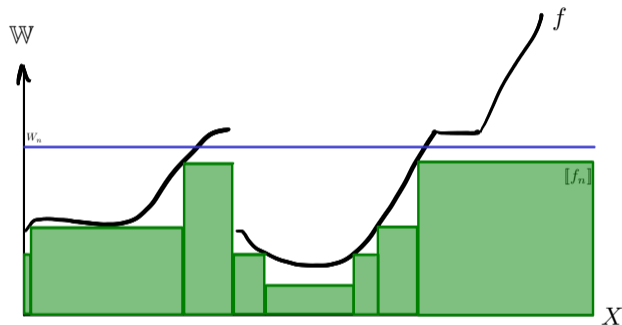
Why?



$\text{approxSimple}_X := \lambda f, \Delta_-, W_-, n. \text{lift let } k = \left\lfloor \frac{W_n}{\Delta_n} \right\rfloor \cdot \Delta_n \text{ in}$

$\left(\begin{array}{l} k + 1, \\ \lambda i. \text{if } i \leq m \text{ then } \lambda x. (i-1) \cdot \Delta_n \leq fx < i \cdot \Delta_n \text{ else } \lambda x. fx = m \cdot \Delta_n, \\ \lambda i. \text{if } i \leq m \text{ then } i \cdot \Delta_n \text{ else } m \cdot \Delta_n \end{array} \right)$

Integrating codes for simple functions



$$\int : \mathbf{GX} \times \mathbf{SimpleCode}_X \rightarrow \mathbb{W}$$

$$\int d\mu(n, E_-, w_-) := \sum_{i=1}^n w_i \cdot \mu E_i$$

Integrating measurable functions

$$\int : \mathbf{GX} \times (X \rightarrow \mathbb{W}) \rightarrow \mathbb{W} \quad \int d\mu f := \sup \{ \int d\mu \varphi \mid \varphi \in \mathbf{SimpleCode}_X, \llbracket \varphi \rrbracket \leq f \}$$
$$= \mathbf{let} \ f_ = \mathbf{approxSimple}_X f (\lambda n. \frac{1}{2^n}) (\lambda n. n) \mathbf{in}$$
$$\lim_{n \rightarrow \infty} \int d\mu f_n$$

A similar construction shows that $\lambda \in \underline{\mathbf{GR}}$.

Giry monad G

Unlike its measure theoretic counterpart, the unrestricted Giry monad over \mathbf{Qbs} is **strong**:

- ▶ Dirac kernel: $\delta_a := \lambda E. \begin{cases} a \in E : 1 \\ a \notin E : 0 \end{cases}$
- ▶ Kock integral: $\oint d\mu k := \lambda E. \int \mu(dx) k(x; E)$

This monad is not suitable as a model for distribution, e.g., it is not commutative, but the counter-example is technical.

Distribution monad \mathbf{D}

Instead, we take those measures and kernels that are laws of the Lebesgue measure:

$$\mathbf{D}X := \{\lambda_\alpha \mid \alpha : \mathbb{R} \rightarrow X\} \quad \mathcal{R}_{\mathbf{D}X} := \{\lambda_r . \lambda_{kr} \mid k : \mathbb{R} \rightarrow (\mathbb{R} \rightarrow X)\}$$

We call such α and k **randomisations**.

Note that $m : \mathbf{D}X \rightarrow \mathbf{G}X$, but may not be a subspace embedding.

- ▶ Dirac lifts along m , via the randomisation $\lambda_x . \lambda_{[\underline{0},1].x, [\underline{0},1]^{\mathbb{C}} \perp}$.
- ▶ To lift the Kock integral, use the fact that there is some measurable isomorphism $\varphi : \mathbb{R} \xrightarrow{\cong} \mathbb{R} \times \mathbb{R}$ whose law is the **2-dimensional** Lebesgue measure λ_2 , i.e., the unique measure on $\mathbb{R} \times \mathbb{R}$ that assigns each rectangle its area:

$$\lambda_2([a,b] \times [c,d]) = (b - a) \cdot (d - c) \quad (a \leq b, c \leq d)$$

\mathbf{D} validates all the properties we listed in Part I, and we use it as our model.

Theorem

For a standard Borel space S :

▶ $\{\mu \in \underline{DS} \mid \|\mu\| = 1\} = \{\mu \in \underline{G'S} \mid \|\mu\| = 1\}$ and similarly for ≤ 1 and $< \infty$.

▶ $\underline{DS} = \{\mu \in \underline{G'S} \mid \mu \text{ is s-finite measure}\}$ and

$\mathcal{R}_{DS} = \{k : \mathbb{R} \rightsquigarrow X \mid k \text{ is s-finite kernel}\}$

where **s-finite** means countable sum of finite distributions/kernels [Staton'16].

Part 1: the **discrete** model (now)

- ▶ Motivation
- ▶ Language of probability and distribution
- ▶ Discrete model
- ▶ Simply-typed probability
- ▶ Dependently-typed probability

Part 2: the **full** model

- ▶ Borel sets and measurable spaces
- ▶ Quasi-Borel spaces
- ▶ Type structure & standard Borel spaces
- ▶ Dependently-typed structure
- ▶ Integration



course page



ask questions on the
Scottish PL Institute
Zulip stream #qbs

Enough! Lunch.